What does the PIN model identify as private information?

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Abstract

We investigate whether the Easley and O'Hara (1987) PIN model's recently documented failure to identify private information arises from the model's inability to describe the data or from the model's reliance on order flows alone. We find that the PIN model mistakenly identifies private information from turnover because it is unable to describe the order flow data. We propose a model that addresses this shortcoming but also depends on order flow alone. We find that the extended model does not perform as well as the Odders-White and Ready (2008) model, which relies on both returns and order flow.

Keywords: Liquidity; Information Asymmetry

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The Probability of Informed Trade (PIN) model, developed in a series of seminal papers including Easley and O'Hara (1987), Easley, Kiefer, O'Hara, and Paperman (1996), and Easley, Kiefer, and O'Hara (1997) has been used extensively in accounting, corporate finance and asset pricing literature as a measure of information asymmetry.[1] The PIN model is based on the notion, originally developed by Glosten and Milgrom (1985), that periods of informed trade can be identified by abnormally large order flow imbalances.[2] Recently, however, several papers have documented PIN anomalies where PINs tend to be at their lowest when information asymmetry should be at its highest (e.g. Aktas, de Bodt, Declerck, and Van Oppens (2007), Benos and Jochec (2007), and Collin-Dufresne and Fos (2014a)).

We address two research questions in this paper. First, we analyze whether PIN mis-identifies private information because the underlying model does not fit the order flow data well. Second, the classic microstructure theories (e.g. Glosten and Milgrom (1985), and Kyle (1985)) suggest that order flow imbalances as well as variables such as prices and spreads are related to the arrival of private information. The PIN model, on the other hand, focuses solely on the response of order flow imbalance to the arrival of private information, ignoring the price response mechanisms that are described in the classic microstructure literature. We therefore analyze the extent to which including the price response mechanism is necessary to empirically identify private information arrival. The answers to these research questions are important because they imply very different agendas for this growing area of research. Specifically, if PIN mis-identifies private information because the model does not fit the order flow data well, then the PIN model could be extended in such a way that it still relies on order flow alone, but no longer mis-identifies private information. On the other hand, if PIN cannot identify private information because it ignores the price response mechanism then a different approach involving variables other than order flow is necessary to generate useful inferences about the arrival of informed trade.

To address these two research questions, we create a variable called the Conditional

[1] A Google scholar search reveals that this series of PIN papers has been cited more than 3,500 times as of this writing. Examples of papers that use PIN in the finance and accounting literature include Duarte, Han, Harford, and Young (2008), Bakke and Whited (2010), Da, Gao, and Jagannathan (2011), and Akins, Ng, and Verdi (2012).

[2] Following the literature we define order flow imbalance as the difference between the number of buyer initiated trades less the number of seller initiated trades. In what follows, we refer to buyer initiated trades as 'buys', seller initiated trades as 'sells', and turnover as the number of buys plus sells.

Probability of an Information Event ($CPIE$). To compute the $CPIE$ implied by the PIN model ($CPIE_{PIN}$), we estimate the PIN model's parameters using an entire year of data, and then use the observed market data (i.e. buys and sells) to estimate the posterior or model-implied probability of an information event for each day in our sample. We then turn to our first question and examine whether observed variation in $CPIE_{PIN}$ is consistent with the theory underlying the PIN model. Under the PIN model, private information is identified solely from the absolute order imbalance. In practice, however, the PIN model may be a poor description of the data and model misspecification can affect the way it actually identifies private information. To test this hypothesis, we regress $CPIE_{PIN}$ for each firm-year on absolute order imbalance, turnover, and their squared terms.

We find that the PIN model primarily identifies information events based on turnover, controlling for absolute order flow imbalance. This is inconsistent with the underlying microstructure assumptions of the model. In regressions of $CPIE_{PIN}$ on absolute order imbalance, turnover, and their squared terms, turnover and turnover squared account for, on average, around 65% of the overall R^2. The identification of information events through turnover becomes more pronounced late in the sample with the increase in both the level and variance of turnover.[3] For example, for the median stock after 2002, the PIN model is essentially equivalent to a naïve model that sets the probability of a private information event equal to one on any day with turnover larger than the annual mean of daily turnover and zero otherwise. Two limitations of the PIN model combine to create this problem. First, under the PIN model, increases in expected turnover can only come about through the arrival of private information. Second, the PIN model's restrictive distributional assumptions make it difficult for the model to match both the mean and the variance of turnover. As a result of these limitations, when confronted with actual data the model mechanically interprets periods of above average turnover as periods of private information arrival.

To show that this conflation of turnover with private information is related to the previous critiques of PIN, we examine an event study similar in spirit to the documented PIN anomalies. For instance, Benos and Jochec (2007) find that PIN is higher after earnings

[3]Duarte and Young (2009) propose an extension of the PIN model that accounts for the positive correlation between buys and sells and thus improves the fit of the model. We show in Internet Appendix A that Duarte and Young (2009) model also performs poorly late in the sample.

announcements than before.[4] In a similar vein, we examine how well the PIN model identifies information events around earnings announcements. In contrast to Benos and Jochec (2007) however, we use $CPIE_{PIN}$ to conduct this event study instead of PIN. There is a large literature (see Bamber, Barron, and Stevens (2011) for a review) that shows that turnover is substantially higher around earnings announcements and typically remains high for a considerable period after the announcement. Since our concern here is the PIN model's ability to separate turnover shocks from information events, earnings announcements provide a good opportunity to examine the model's performance and allows us to connect our results with those in previous studies. As in our full-sample regressions, our event study shows that $CPIE_{PIN}$ is higher after announcements simply due to the higher levels of turnover in the post-announcement periods.

This mechanical conflation of increases in turnover with the arrival of private information in the PIN model is a problem because it implies that the most popular measure of private information in the literature, PIN, does not actually capture its variable of interest. There is no theoretical reason why turnover should be mechanically associated with the arrival of private information, once we control for order imbalance. On one hand, trading by informed traders may increase turnover. On the other hand, liquidity traders may postpone trading when the arrival of private information is likely leading to a negative relation between turnover and private information (e.g. Chae (2005)). Moreover, a model that naively associates turnover with private information arrival ignores the fact that turnover varies for reasons unrelated to private information. For instance, turnover can increase with disagreement (e.g. Kandel and Pearson (1995), and Banerjee and Kremer (2010)). Turnover is also subject to calendar effects because traders coordinate trade on certain days to reduce trading costs (Admati and Pfleiderer (1988)). Furthermore, turnover can vary due to portfolio rebalancing (Lo and Wang (2000)) and taxation reasons (Lakonishok and Smidt (1986)).[5] Hence, the PIN model (and the PIN measure) groups all sources of variation in turnover (e.g. disagreement, calen-

[4]In addition, Aktas, de Bodt, Declerck, and Van Oppens (2007) find that PIN is higher after merger announcements than before.

[5]The literature also suggests that turnover after earnings announcements can remain high for many reasons unrelated to the arrival of private information. For instance, traditionally, the literature attributes high turnover after announcements to disagreement (e.g. Bamber, Barron, and Stevens (2011)). Karpoff (1986) suggests that high turnover after earnings announcements may also be due to divergent prior expectations, while Frazzini and Lamont (2007) attribute to small investors' lack of attention.

dar effects, portfolio rebalancing, taxation, etc.) under the umbrella of private information arrival.

Having demonstrated that the PIN model essentially treats all shocks to turnover as private information because it fits the data so poorly, we turn to our second research question. Namely we analyze the extent to which a model that includes the price response mechanism generates better inferences about the arrival of informed trade than a model based on order flow alone. To do so, we compare an extension of the PIN model (the EPIN model) with the model developed by Odders-White and Ready (2008) (the OWR model). The EPIN model is based on the same information structure as the PIN model. The key difference is that the EPIN model fixes the PIN model's mechanical conflation of turnover and private information arrival. In contrast to the PIN and EPIN models, the OWR model is based on Kyle (1985) and uses intraday as well as overnight returns, along with order imbalance, to identify private information events.

We use the EPIN and OWR $CPIE$s ($CPIE_{EPIN}$ and $CPIE_{OWR}$) to compare the models in three different ways.[6] First, under the assumption that private information should arrive prior to earnings announcements, rather than after the announcement, we expect that if a model correctly identifies informed trade, its $CPIE$ will increase prior to the announcement. We also anticipate that informed trading, and hence $CPIE$s, will decline rapidly after the announcement, when investors have the same (now public) information.[7] Second, we follow Cohen, Malloy, and Pomorski (2012) and identify instances of opportunistic insider trades. If either of the models can successfully detect opportunistic insider trading, then its $CPIE$ should increase around these trades. Third, it has long been recognized in the literature (e.g. Hasbrouck (1988, 1991a,b)) that non-information related price changes (e.g. dealer inventory control) should be subsequently reversed, while information related trades should not. Therefore, if a model correctly identifies the arrival of private information, we expect that increases in its $CPIE$ should be associated with smaller future price reversals. Each of

[6] While the PIN and EPIN models allow for a calculation of the probability of informed trade, the OWR model does not. However, all three models have a parameter that controls the unconditional probability of an information event on a given day (α) and allow for the calculation of $CPIE$.

[7] There is considerable evidence suggesting the possibility of high asymmetric information prior to important announcements. See for example Brooks (1996), Meulbroek (1992) Christophe, Ferri, and Angel (2004), Amin and Lee (1997), Frazzini and Lamont (2007), and Hendershott, Livdan, and Schurhoff (2014).

these three methods of model comparison has its own unique limitations.[8] However, if all of these methods point to the same conclusions, it seems unlikely that our overall interpretation would be biased due to the limitations of any specific method.

In answer to our second question, we find that the OWR model performs better than the EPIN model in all three tests. Specifically, we find that the $CPIE_{OWR}$ increases before earnings announcements and decreases rapidly after announcements, while $CPIE_{EPIN}$ decreases before announcements. $CPIE_{OWR}$ successfully predicts opportunistic insider trading and is strongly negatively associated with price reversals. In contrast, $CPIE_{EPIN}$ is only weakly associated with opportunistic insider and price reversals.

We contribute to the literature because we show that private information measures based only on order flow (e.g. PIN) perform much worse than those that include the price response mechanism, for instance the OWR's α. The classic microstructure theories (e.g. Glosten and Milgrom (1985), and Kyle (1985)) describe a price response mechanism relating returns to the arrival of private information. Hence it is not surprising that a model that identifies the arrival of private information solely from order flow imbalance has worse performance than a model based on returns and order flows. However, it is perhaps surprising that the OWR model performs so much better than the EPIN model in all of our tests. This suggests that order flow, however well modeled, is insufficient to be the sole source of inferences about private information arrival. Therefore, despite the literature's strong interest in proxies of private information based on order flow alone (e.g. Easley, Kiefer, O'Hara, and Paperman (1996), Easley, Kiefer, and O'Hara (1997), and Duarte and Young (2009)) future research aimed at building measures of informed trade should also focus on variables such as prices and spreads as the classic theory suggests.

Our paper is also related to a growing literature that analyzes the extent to which PIN actually captures information asymmetry. Duarte and Young (2009) and Gan, Wei, and Johnstone (2014) show that the PIN model does not fit the order flow data well. We take these results one step further and show that *because* of this poor fit the PIN model mis-identifies the variable of interest—private information—from turnover. In addition, we extend the

[8]For instance, it is possible that, for some reason, private information is more prevalent after important announcements than before.

PIN model to correct the mechanical conflation of turnover and private information arrival. This allows us to address whether order flow alone can capture private information arrival or whether we must incorporate the price response mechanism as in the OWR model. Many of the papers analyzing the PIN measure estimate PINs around events and test whether PIN is higher before rather than after an announcement. These studies in general document that PIN is higher after announcements than before (i.e. PIN anomalies). For instance, Collin-Dufresne and Fos (2014a) find that PIN and other adverse selection measures are lower when Schedule 13D filers trade.[9] Easley, Engle, O'Hara, and Wu (2008) critique this line of research, noting that PIN is a stock characteristic rather than a measure of the extent to which private information is present in a given calendar time period.[10] To address this critique, Easley, Engle, O'Hara, and Wu (2008) develop an extension of the original model in which PIN is time-varying, and in a paper contemporaneous to ours, Brennan, Huh, and Subrahmanyam (2015) use conditional probabilities similar to $CPIE_{PIN}$. We contribute to this literature in two ways. First, our results indicate that these previously identified PIN anomalies are at least partially related to the strong connection between $CPIE_{PIN}$ and turnover that we document. Second, we show that event studies that use daily measures of private information (e.g. Easley, Engle, O'Hara, and Wu (2008)) can be misleading if variation in these measures around event announcements is due to variables not necessarily related to information asymmetry. For instance, Brennan, Huh, and Subrahmanyam (2015) interpret the fact that their $CPIE_{PIN}$ measures are higher after earnings announcements than before as evidence of informed trading. We show that $CPIE_{PIN}$ is naively related to turnover. This suggests that the findings in Brennan, Huh, and Subrahmanyam (2015) can simply be attributed to the fact that turnover is typically much higher after earnings announcements.

[9]Collin-Dufresne and Fos (2014b) partially attribute this finding to informed traders disguising their trades in periods of high liquidity or timing their trades such that market movements conceal the nature of their information. Our findings cannot speak to this possibility, instead we show that the PIN model mechanically attributes all sources of variation in turnover to the arrival of private information.

[10]Easley, Lopez de Prado, and O'Hara (2012) develop the volume-synchronized probability of informed trading or $VPIN$. We do not consider $VPIN$ in this paper because, as Easley, Lopez de Prado, and O'Hara (2012) point out, $VPIN$ is a measure of order flow toxicity at high frequencies rather than a stock characteristic that measures adverse selection at lower frequencies as PIN is widely used in the finance and accounting literature. Moreover, Andersen and Bondarenko (2014) provide detailed critique of the $VPIN$ measure.

The remainder of the paper is as follows. Section 1 outlines the data we use for our empirical results. Section 2 shows that the PIN model mechanically associates variation in turnover with the arrival of private information. Section 3 extends the PIN model to deal with this shortcoming and compares a model based on order flow imbalance alone (EPIN) with a model that identifies private information from both returns and order flow (OWR). Section 4 concludes.

1 Data

To estimate the PIN, EPIN, and OWR models, we collect trades and quotes data for all NYSE stocks between 1993 and 2012 from the NYSE TAQ database. We require that the stocks in our sample have only one issue (i.e. one `PERMNO`), are common stocks (share code 10 or 11), are listed on the NYSE (exchange code 1), and have at least 200 days worth of non-missing observations for the year. Our sample contains 1,060 stocks per year on average. Despite our sample selection criteria, about 36% (25%) of the stocks in our sample are in the top (bottom) three Fama-French size deciles. For each stock in the sample, we classify each day's trades as either buys or sells, following the Lee and Ready (1991) algorithm. In our analysis, we define turnover as the sum of daily buys and sells. Internet Appendix B describes the computation of the number of buys and sells.

We estimate both the PIN and EPIN models using only the daily number of buys and sells ($B_{i,t}$ and $S_{i,t}$). The OWR model, however, also requires intraday and overnight returns as well as order imbalances. Following Odders-White and Ready (2008) we compute the intraday return at day t as the volume-weighted average price (VWAP) at t minus the opening quote midpoint at t plus dividends at time t, all divided by the opening quote midpoint at time t.[11] We compute the overnight return at t as the opening quote midpoint at $t+1$ minus the VWAP at t, all divided by the opening quote midpoint at t. The total return, or sum of the intraday and overnight returns is the open-to-open return from t to $t+1$. We compute order imbalance (y_e) as the daily share volume of buys minus the share

[11] The opening quote midpoint is not available in TAQ in many instances. When the opening quote midpoint is not available, we use the matched quote of the first trade in the day as a proxy for the opening quote.

volume of sells, divided by the total share volume. We follow Odders-White and Ready and remove systematic effects from returns to obtain measures of unexpected overnight and intraday returns ($r_{o,i,t}$ and $r_{d,i,t}$). See Internet Appendix B for details.

Like Odders-White and Ready (2008), we remove days around unusual distributions or large dividends, as well as CUSIP or ticker changes. We also drop days for which we are missing overnight returns ($r_{o,i,t}$), intraday returns ($r_{d,i,t}$), order imbalance (y_e), buys (B), or sells (S). Our empirical procedures follow those of Odders-White and Ready with two exceptions. First, OWR estimate y_e as the idiosyncratic component of net order flow divided by shares outstanding. We do not follow the same procedure as OWR in defining y_e because we find that estimating y_e as we do results in less noisy estimates. Specifically, we find that y_e defined as shares bought minus shares sold divided by shares outstanding, as in Odders-White and Ready (2008), suffers from scale effects late in the sample, when order flow is several orders of magnitude larger than shares outstanding. Second, Odders-White and Ready remove a whole trading year of data surrounding distribution events, but we only remove one trading week [-2,+2] around these events.

For the event study portion of our analysis, we examine earnings announcements. Our sample of earnings announcements includes all CRSP/COMPUSTAT firms listed in NYSE between 1995–2009 for which we have exact timestamps collected from press releases in Factiva which fall within a [-1,0] window relative to COMPUSTAT earnings announcement dates following Dong, Li, Ramesh, and Shen (2015). Because we have exact timestamps for the earnings announcements, we can cleanly separate between the pre and post event periods, thus avoiding ambiguity about when exactly the information becomes public. To avoid any confusion with respect of the timing of the events in the OWR model, we remove all announcements occurring on non-trading days. Our final sample includes 21,979 earnings announcements.

We also examine a sample of opportunistic insider trades, as defined in Cohen, Malloy, and Pomorski (2012), from the Thomson Reuters' database of insider trades. In order to classify a trader as opportunistic or routine, we require three years of consecutive insider trades. We classify a trader as routine if she places a trade in the same calendar month for at least three years. All non-routine traders' trades are classified as opportunistic. Cohen, Malloy, and Pomorski (2012) show that opportunistic insider trades predict abnormal returns, information events, and regulator actions, which is consistent with the presence of private information. Our event sample includes 32,676 opportunistic insider trades.

Table 1 contains summary statistics of all the variables used to estimate the models. Panel A gives summary statistics of our entire sample, Panel B displays the summary statistics for the days of earnings announcements, and Panel C displays the summary statistics for opportunistic insider trading days.

2 Why does PIN fail?

This section addresses whether PIN mis-identifies private information because the underlying model does not fit the data well. Section 2.1 briefly describes the PIN model and $CPIE_{PIN}$. Section 2.2 shows the results of regressions of $CPIE_{PIN}$ on absolute order imbalance and turnover. Section 2.3 shows how $CPIE_{PIN}$ varies around earnings announcements. The results in Sections 2.2 and 2.3 show that the PIN model identifies the arrival of private information from increases in turnover.

2.1 Description of the PIN model

The Easley, Kiefer, O'Hara, and Paperman (1996) PIN model posits the existence of a liquidity provider who receives buy and sell orders from both informed traders and uninformed traders. At the beginning of each day, the informed traders receive a private signal with probability α. If the private signal is positive (which occurs with probability δ), buy orders from informed and uninformed traders arrive following a Poisson distribution with intensity $\mu + \epsilon_B$, while sell orders come only from the uninformed traders and arrive with intensity ϵ_S. If the private signal is negative (with probability $1 - \delta$), sell orders from informed and uninformed traders arrive following a Poisson distribution with intensity $\mu + \epsilon_S$, while buy orders come only from the uninformed traders and arrive with intensity ϵ_B. If the informed traders receive no private signal, they do not trade; thus, all buy and sell orders come from the uninformed traders and arrive with intensity ϵ_B and ϵ_S, respectively. Fig. 1 shows a tree diagram of this model. The difference in arrival rates captures the intuition that on days with positive private information, the arrival rate of buy orders increases over and above the normal rate of noise trading because informed traders enter the market to place buy orders. Similarly, the arrival rate of sell orders rises when the informed traders seek to sell based on their negative private signals. Therefore, the PIN model identifies the arrival of private information through increases in the absolute value of the order imbalance.

The model also ties variations in turnover to the arrival of private information. Specifically, let the indicator $I_{i,t}$ take the value of one if an information event occurs for stock i on day t, and zero otherwise. Note that under the model the number of buys plus sells (turnover) is distributed as a Poisson random variable with intensity:

$$\lambda(I_{i,t}) = \begin{cases} \epsilon_B + \epsilon_S & \text{when } I_{i,t} = 0 \\ \epsilon_B + \epsilon_S + \mu & \text{when } I_{i,t} = 1 \end{cases} \tag{1}$$

Thus, under the PIN model, private information is *necessarily* the cause of any variation in expected daily turnover.

To formalize the concept of $CPIE_{PIN}$, let $B_{i,t}$ ($S_{i,t}$) represent the number of buys (sells) for stock i on day t and $\Theta_{PIN,i} = (\alpha_i, \mu_i, \epsilon_{B_i}, \epsilon_{S_i}, \delta_i)$ represent the vector of the PIN model parameters for stock i. Let $D_{PIN,i,t} = [\Theta_{PIN,i}, B_{i,t}, S_{i,t}]$. The likelihood function of the Easley, Kiefer, O'Hara, and Paperman (1996) model is $\prod_{t=1}^{T} L(D_{PIN,i,t})$, where $L(D_{PIN,i,t})$ is equal to the likelihood of observing $B_{i,t}$ and $S_{i,t}$ on a day without private information ($L_{NI}(D_{PIN,i,t})$) added to the likelihood of $B_{i,t}$ and $S_{i,t}$ on a day with positive information ($L_{I^+}(D_{PIN,i,t})$) and to the likelihood of $B_{i,t}$ and $S_{i,t}$ on a day with negative information ($L_{I^-}(D_{PIN,i,t})$). Each of the likelihood functions ($L_{NI}(D_{PIN,i,t})$, $L_{I^+}(D_{PIN,i,t})$ and $L_{I^-}(D_{PIN,i,t})$) corresponds to a node of the tree in Fig. 1. See Internet Appendix C for details.

Using the PIN model, for each stock-day, we compute the probability of an information event conditional both on the model parameters and on the observed total number of buys and sells. For the PIN model, we compute $CPIE_{PIN,i,t} = P\left[I_{i,t} = 1 | D_{PIN,i,t}\right]$. This probability is given by $(L_{I^-}(D_{PIN,i,t}) + L_{I^+}(D_{PIN,i,t}))/L(D_{PIN,i,t})$. $CPIE_{PIN,i,t}$ represents the econometrician's posterior probability of an information event given the data observed on that day, and the underlying model parameters.

Note that if we condition down with respect to the data, $CPIE_{PIN,i,t}$ reduces to the model's unconditional probability of information events (α_i). The unconditional probability represents the econometrician's beliefs about the likelihood of an information event before seeing any actual orders or trades. In the absence of buy and sell data, an econometrician would assign a probability α_i to an information event for stock i on day t, where $\alpha_i = E[CPIE_{PIN,i,t}]$ and the expectation is taken with respect to the joint distribution of $B_{i,t}$ and $S_{i,t}$. The PIN of a stock, defined as $\frac{\alpha\mu}{\alpha\mu+\epsilon_B+\epsilon_S}$, is the unconditional probability that any given trade is initiated by an informed trader. $CPIE$ and PIN are linked via the unconditional probability of an information event, α.

We estimate the PIN model numerically via maximum likelihood for every firm-year in our sample. The estimation procedure is similar to that used in Duarte and Young (2009). The parameter estimates are used for computing $CPIE_{PIN}$ in Sections 2.2 and 2.3. Internet Appendix C provides details about the maximum likelihood procedure and the calculation of $CPIE_{PIN}$.

Table 2 contains summary statistics for the parameter estimates of the PIN model. Table 2 also contains summary statistics of the cross-sectional sample means and standard deviations of $CPIE_{PIN}$. The results in Table 2 show that the mean $CPIE_{PIN}$ behaves exactly like α. Hence, changes in $CPIE_{PIN}$ and changes in the estimated α are analogous. Fig. 2 Panel A shows how the distribution of α changes over time. Interestingly, the PIN model α increases over time, with the median PIN α rising from about 30% in 1993 to 50% in 2012.[12] Panel B of Fig. 2 plots the time series of PIN. Note that PIN decreases over time in spite of the fact that α increases. This happens because, according to the PIN model, the intensity of noise trading is increasing over time while the intensity of informed trading remains relatively flat as shown in Panel C of Fig. 2. It is important to note, however, that the time series patterns of the model parameters in Fig. 2 have no implications for how the PIN model identifies private information.

We also estimate the parameter vectors $\Theta_{PIN,i}$ in the period $t \in [-312, -60]$ before an earnings announcement. These parameter estimates are used to compute the $CPIE$s in Section 2.3. The summary statistics of the parameter estimates for the event studies are qualitatively similar to those in Table 2 and in Figure 2.

[12]The increase in our estimated PIN model α parameters is somewhat larger than that in Brennan, Huh, and Subrahmanyam (2015). This small difference arises because Brennan, Huh, and Subrahmanyam (2015) have a larger number of stocks per year due to the fact that we apply sample filters similar to those in Odders-White and Ready (2008). In fact, without these filters, the increase in our estimated PIN model α parameters from 1993 to 2012 is comparable to that in Brennan, Huh, and Subrahmanyam (2015).

2.2 How does the PIN model identify private information?

This section analyzes how the PIN model actually identifies private information. In theory, the PIN model identifies information events from changes in the absolute order flow imbalance. Empirically, however, the PIN model may produce such a poor description of the order flow data that the model actually mis-identifies the variable of interest — private information. To analyze how the PIN model identifies private information in practice, we regress $CPIEs$ on absolute order imbalance and turnover in Section 2.2.1. The results of these regressions show that on average 65% of the variation in $CPIE_{PIN}$ is explained by turnover instead of absolute order imbalance. The intuition for this failure of the PIN model can be clearly seen in the scatter plot of buys and sells for Exxon-Mobil in Section 2.2.2. This scatter plot shows that the model mechanically identifies the arrival of private information from turnover. In fact, the PIN model essentially assigns probability one to the arrival of private information on any day when turnover is above the average daily turnover in the year and zero otherwise. As a result, the PIN model naively groups all sources of variation in turnover (e.g. disagreement, calendar effects, portfolio rebalancing, taxation, etc.) under the umbrella of private information arrival. In Section 2.2.3, we show that this naive identification of private information happens not only for Exxon-Mobil but also for the majority of the stocks in our sample following the increase in turnover in the early 2000s.

Given the strong connection between $CPIE$s and the unconditional probability of information arrival (α), our results in this section call into question the use of PIN as proxy for private information. While there are other parameters in the model (i.e. μ, ϵ_B and ϵ_S), these parameters are jointly identified with α. Hence it seems extremely unlikely that in the joint identification of the model parameters, biases in the other parameters 'correct' the biases in α in such a way that PIN is 'rescued' as a reasonable proxy for private information. Thus, while our $CPIE$ results do not speak directly to μ, ϵ_B and ϵ_S, they still call into question PIN as a measure of private information.

2.2.1 Regression Tests

Since there are many moments that the PIN model can fail to match, there are many tests that might reject the PIN model (e.g. Duarte and Young (2009)). Our regression tests are not designed to analyze whether the PIN model matches particular moments in the data but instead are focused on how the PIN model identifies the fundamental variable of interest—private information. Specifically, our analysis is anchored around the regression $CPIE_{PIN} = \alpha + \beta_0|B-S| + \beta_1|B-S|^2 + \beta_2 turn + \beta_3 turn^2 + \varepsilon$. Since $CPIE_{PIN}$ is a direct measure of private information according to the PIN model, this regression reveals how the PIN model actually identifies private information.

To formally show that the PIN model identifies private information from turnover instead of order flow, we compare the results from regressions with data created by simulating the PIN model to results from regressions with real data. To create the simulated data, we first estimate the parameters of the PIN model for each firm-year in our sample. Then, for each firm-year, we generate 1,000 artificial firm-years' worth of data (i.e. $B_{i,t}$ and $S_{i,t}$) using the estimated parameters. We then compute the $CPIE_{PIN,i,t}$ for each trading day in a simulated trading year and regress these $CPIE$s absolute order flow imbalance and turnover. The results of the regressions using simulated data are useful because they reveal how the PIN model is intended to identify private information arrival and also allow us to build empirical distributions of the R^2s of the regressions of $CPIE$s on order imbalance and turnover under the null hypothesis that the PIN model correctly describes the order flow data.

Panel A of Table 3 presents the results of yearly multivariate regressions of $CPIE_{PIN}$ on absolute order flow imbalance $|B-S|$ and $|B-S|^2$. We add squared terms to these regressions to account for nonlinearities in the relationship between $CPIE_{PIN}$ and $|B-S|$. We average the simulated results for each `PERMNO`-Year and report in Panel A of Table 3 the median coefficient estimates and t-statistics. The coefficients are standardized so they represent the increase in $CPIE_{PIN}$ due to a one standard deviation increase in the corresponding independent variable. We also report the average of the median, the 5^{th}, and the 95^{th} percentiles of the empirical distribution of R^2s of these regressions generated by the 1,000 simulations. In general, the coefficients are highly statistically significant and the R^2s are high. This is consistent with intuition that if the model were literally true, the absolute order imbalance could be used to infer the arrival of private information.

The columns of Table 3 labeled as '$R^2_{inc.}$' include statistics on the increase in the R^2 that is due to the inclusion of turnover ($turn$) and turnover squared ($turn^2$) in the regressions. Specifically, $R^2_{inc.}$ is equal to the difference between the R^2 of the extended regression model with turnover terms and the R^2 of a regression that includes only order imbalance terms. We report the average of the median, the 5^{th}, and the 95^{th} percentiles of the $R^2_{inc.}$s of these regressions across the 1,000 simulations. The incremental increase in R^2s are relatively low, with an average value of around 10%, which implies that, under the model's data generating process, turnover has only modest incremental power in explaining $CPIE_{PIN}$. The picture that emerges from these regressions is that if the PIN model were a perfectly accurate representation of trading activity, $CPIE_{PIN}$ would be determined solely by the order flow imbalance on each day.

Panel B of Table 3 reports regression results for the real rather than simulated data. With the real data, the picture is very different. The R^2s of the regressions of $CPIE_{PIN}$ on $|B - S|$ and $|B - S|^2$ are much smaller than those in the simulations. On the other hand, the incremental R^2s from turnover are much higher than those in Panel A. The incremental R^2 also increases over time with a value of about 36% in 1993, to nearly 46% in 2012. This implies that turnover and turnover squared explain a much larger degree of variation in $CPIE_{PIN}$ than order imbalance. In fact, the average ratio of the median R^2s, $R^2_{inc.}/(R^2 + R^2_{inc.})$, is about 65%. The difference arises because, in the real data, absolute order flow and turnover are only weakly correlated. For instance, large absolute order flow imbalances are possible when turnover is below average, and vice versa. Under the PIN model, however, the two are highly correlated.

We test the hypothesis that $R^2_{inc.}$s in the actual data are consistent with those generated under the PIN model. Panel B reports the average p-value (the probability of observing an $R^2_{inc.}$ in the simulations at least as large as what we observe in the data) across all stocks, and the frequency that we reject the null at the 5% level implied by the distribution of simulated $R^2_{inc.}$s. The PIN model is rejected in about 89% of the stock-years in our sample, and there is on average less than a 7% chance of the PIN model generating $R^2_{inc.}$s as high as what we see in the data.

The results in Table 3 indicate that the PIN model identifies private information from increases in turnover, as opposed to changes in order imbalances for the majority of the sample. These findings are inconsistent with the microstructure assumptions of the PIN model—controlling for order imbalance there should be no room for turnover in explaining private information arrival.

2.2.2 Exxon-Mobil Scatter Plots

To understand the intuition behind the results in Table 3, consider the scatter plot of real and simulated order flow data for Exxon-Mobil in Fig. 3. Panels A and B plot simulated and real order flow for Exxon-Mobil in 1993 and 2012 respectively, with buys on the horizontal axis and sells on the vertical axis. Real data are marked as +, and simulated data as transparent dots. The real data are shaded according to the $CPIE$, with darker points (+ **magenta**) representing low and lighter points (+ cyan) high $CPIE$s. Panels C and D plot the CPIE$_{PIN}$ as function of turnover. The vertical lines in these panels represent the annual mean of daily turnover.

Panel A of Fig. 3 illustrates the central intuition behind the PIN model. The simulated data comprise three types of days, which create three distinct clusters. Two of the clusters are made up of days characterized by relatively large order flow imbalance, with a large number of sells (buys) and relatively few buys (sells). The third group of days has relatively low numbers of buys and sells because there is no private information arrival. Generalizing from this figure, days with large order flow imbalances correspond to informed traders entering the market in the PIN model.

The real data, on the other hand, show no distinct clusters in Panel A, and in Panel B of Fig. 3 the PIN model's three clusters barely overlap with even a small portion of the data. This implies that the model cannot account for existence of the majority of the daily observations of order flow for Exxon-Mobil in 2012. In essence, the model classifies almost all daily observations as extreme outliers. The intuition for this is that the PIN model assumes that order flow is distributed as a mixture of three bivariate Poisson random variables (i.e. the three clusters in Panels A and B). The mean and the variance of a Poisson random variable are equal and, as a consequence, the Poisson mixtures behind the PIN model cannot accommodate the high level and volatility of turnover that we observe, especially in the later part of the sample.

Panels A and B also plot a line that separates the scatter plots in two regions. All the observations below (above) these lines have turnover below (above) the annual mean of daily turnover. These lines along with the $CPIE$ color scheme for the observed data suggest that the PIN model is mechanically identifying private information from turnover. To clarify this mechanical identification, Panels C and D plot $CPIE_{PIN}$ as function of turnover. Panels C and D show that the PIN model essentially classifies days with above average turnover as private information days (i.e. $CPIE_{PIN}$ equal to one) and days with below average turnover as days without private information (i.e. $CPIE_{PIN}$ equal to zero). The reason for this mechanical conflation of turnover with private information arrival is that under the PIN model expected turnover can only vary because of the arrival of private information (see Equation 1). Hence the poor fit to the turnover data along with the connection between turnover and arrival of private information in the PIN model causes the model to mechanically identify shocks to turnover as due to the arrival of private information.

Fig. 3 also emphasizes the mechanical nature of the relation between $CPIE_{PIN}$ and turnover. In 2012, the PIN model identifies almost all days with higher than average turnover as days with private information events. Note that this identification does not necessarily relate to the possibility, suggested by Collin-Dufresne and Fos (2014b), that informed traders sometimes choose to trade on days with high liquidity or turnover. Naturally, it is possible that informed traders do in fact trade on some days with high turnover. However, the point here is that the PIN model identifies essentially *all* days with above average turnover as information events.

2.2.3 $CPIE_{Naive}$

Fig. 3 shows the PIN model's naive identification of private information events for one stock, and in this section we show that this is not an isolated example. In fact, the problem is widespread. To quantify how often the PIN model classifies information events as simple function of turnover we define

$$CPIE_{Naive,i,t} = \begin{cases} 0, & \text{if } turn_{i,t} < \overline{turn}_i \\ 1, & \text{if } turn_{i,t} \geq \overline{turn}_i \end{cases} \tag{2}$$

That is, $CPIE_{Naive,i,t}$ is a dummy variable equal to one when turnover for stock i on day t ($turn_{i,t}$) is larger than or equal to the annual average of daily turnover of stock i ($\overline{turn}_i$) and zero otherwise. To our knowledge there is no paper in the literature that proposes identifying private information in similar manner.[13] It is clear, however, from Panel D of Fig. 3 that the PIN model essentially identifies the arrival of private information for Exxon-Mobil in 2012 according to this rule. We use $CPIE_{Naive}$ to gauge the extent to which the PIN model conflates the arrival of private information with turnover. Specifically, Panel A of Fig. 4 shows the distribution of the fraction of days for which $CPIE_{PIN}$ is identical to $CPIE_{Naive}$ ($|CPIE_{PIN} - CPIE_{Naive}| < 10^{-10}$). $CPIE_{PIN}$ and $CPIE_{Naive}$ are identical for about 85% of the annual observations for the median stock since 2002.

Another way to gauge the extent to which the PIN model breaks down later in our sample period is to count the number of days that the PIN model classifies as outliers. Panel B of Fig. 4 shows the fraction of days for the median stock-year which the PIN model classifies as "outliers" (likelihoods smaller than 10^{-10}). According to the PIN model, for the median stock about 60% (90%) of the annual observations are classified as outliers in 2005 (2010).[14]

Figs. 3 and 4 also give the intuition for why the median PIN α increases over time in Fig. 2. To see this, recall that α is the unconditional expected value of $CPIE_{PIN}$. Therefore, as we observe more $CPIE_{PIN}$ values approaching one, the estimated PIN α must increase. In fact, the median PIN α becomes close to 50% later in the sample which consistent with the fact that the PIN model assigns a $CPIE_{PIN}$ equal to one (zero) to days with turnover above (below) the average.

2.3 Relating PIN anomalies to turnover

The previous section shows that the PIN model often identifies private information from turnover. The question remains, however, whether this is merely an inconsequential specification issue or whether this changes the interpretation of results in the existing literature (e.g. Aktas, de Bodt, Declerck, and Van Oppens (2007), Benos and Jochec (2007), Brennan, Huh, and Subrahmanyam (2015), and Easley, Engle, O'Hara, and Wu (2008)). To

[13] Stickel and Verrecchia (1994) propose identifying information arrival in general with a similar measure, but not private information in particular.

[14] O'Hara, Yao, and Ye (2014) find that high-frequency trading is associated with an increase in the use of odd lot trades, which do not appear in the TAQ database. Therefore, estimates of the PIN model parameters computed using recent TAQ data may be systematically biased. More broadly, Fig. 4 indicates that even if the PIN model are estimated using data that includes odd lot trades, the model will still be badly misspecified late in the sample.

address this, we examine how well the PIN model identifies information events around earnings announcements. Turnover is typically much higher around earnings announcements (e.g. Bamber, Barron, and Stevens (2011)) hence earnings announcements provide a good laboratory to examine this question.

Unlike a standard event study, we focus on movements in $CPIE$ rather than price movements. For each model, we examine the period $t \in [-20, 20]$ around the event. To do so, we estimate the parameter vector $\Theta_{PIN,i}$ in the period $t \in [-312, -60]$ before the event and then compute the daily $CPIE$s for the period $t \in [-20, 20]$ surrounding the announcement. Prior studies estimate the parameters of the model in various windows around an event in order to compute the PIN. Our procedure is different in that we estimate the parameters of the model one year prior to the event and then employ the estimated parameters as if we were an econometrician observing the market data (i.e. buys and sells) and attempting to infer whether an information event occurred. Table 1 Panel B presents summary statistics for order imbalance, intraday returns, overnight returns, number of buys, and the number of sells for earnings announcement days ($t = 0$).

Panel A of Fig. 5 shows the average $CPIE_{PIN}$ in event time for our sample of earnings announcements. The graph shows that, under the PIN model, the probability of an information event increases prior to the event, starting below 55% 20 days before the announcement and peaking above 80% on the day after the announcement. The rise in the probability of an information event prior to the announcement could be consistent with a world where informed traders generate signals about earnings and trade on this information before earnings are announced to the public. However, $CPIE_{PIN}$ is also higher *after* the actual earnings become public information.

Panels B and C of Fig. 5 shed light on the features of the data that produce the observed pattern in the average $CPIE_{PIN}$ in Panel A. Panel B shows the average predictions from OLS regressions of $CPIE_{PIN}$ on order imbalance and absolute order imbalance squared across all of the stocks in the event study sample. The solid line indicates that order imbalance explains only a small fraction of the variation in $CPIE_{PIN}$ within the event window. Panel C shows the average predictions from regressions of $CPIE_{PIN}$ on turnover and turnover squared. The solid line indicates that the variation in $CPIE_{PIN}$ around earnings announcements is explained almost entirely by turnover. The intuition follows directly from the results in Section 2.2, which shows that $CPIE_{PIN}$ is mechanically driven by turnover increases. The higher post-event turnover levels are enough to keep $CPIE_{PIN}$ above its pre-event mean for a substantial period.

To formalize the intuition behind Panels B and C of Fig. 5, we run regressions similar to those in Table 3 using our event sample. Specifically, we run regressions of $CPIE_{PIN}$ on absolute value of order imbalance and its squared term during the event window [-20,+20]. The results of these regressions (see Table 4) indicate that absolute order imbalance explains little of the variation in $CPIE_{PIN}$ in the event window while turnover explains most of the variation in $CPIE_{PIN}$. In fact, Table 4 shows that for the median stock, adding turnover and turnover squared to these regressions nearly quadruples the R^2s.

The event study results suggest that the variation in PIN around events documented in the literature is partially related to variation in α that is mechanically driven by turnover, rather than order imbalance. For instance, Benos and Jochec (2007) show that PIN increases after earnings announcements, while Aktas, de Bodt, Declerck, and Van Oppens (2007) show that PIN increases after M&A target announcements due to increases in both μ and α. Therefore, our evidence suggests that these PIN results are at least partially explained by the fact that the PIN model attributes increases in turnover to private information.

Turnover around earnings announcement can vary for many reasons unrelated to the arrival of private information. Traditionally the literature has attributed high turnover after announcements to disagreement (e.g. Bamber, Barron, and Stevens (2011)). Karpoff (1986) suggests that high turnover after earnings announcements may also be due to divergent prior expectations, while Frazzini and Lamont (2007) attributes high turnover to small investors' lack of attention. None of these studies suggest that the higher turnover around announcements is necessarily the result of increased informed trade, per se. Even the PIN model suggests that once we control for order imbalance, turnover should have little power to identify informed trade.

Another important implication of these results for the literature is that event studies based on daily measures of private information, like $CPIE_{PIN}$ (e.g. Easley, Engle, O'Hara, and Wu (2008) and Brennan, Huh, and Subrahmanyam (2015)) can also be misleading. To see this point consider the results in Panel A of Fig. 5. It may appear at first glance that the results in Panel A of Fig. 5 suggest that the PIN model identifies private information in a sensible way since $CPIE_{PIN}$ increases dramatically from 55% before the announcement to over 75% on the day of the announcement then falls after the announcement, albeit over a period of weeks. However, the decomposition of the $CPIE$s in Panels B and C of Fig. 5 points to a different interpretation, namely that the dramatic increase in $CPIE$ around the event is actually result of variation in turnover, which may be unrelated to the arrival of private information as we point out above.

3 Does order flow alone reveal private information?

The previous section shows that the PIN model mis-identifies private information arrival from increases in turnover. However, it could be that net order flow itself is such a poor indicator of private information that no model based on order flow alone is capable of identifying informed trade (e.g. Back, Crotty, and Li (2014) and Kim and Stoll (2014)). This section gauges the extent to which a model of order flows and price responses generates better inferences about the arrival of informed trade than a model based on order flow alone. To do so, we first propose an extension of the PIN model (the EPIN model) that removes the mechanical conflation of turnover and arrival of private information that plagues the PIN model. We then compare the OWR model, which infers the arrival of private information from returns and order flow, with the EPIN model, which is solely based on order flow. Section 3.1 presents the EPIN model. Section 3.2 describes the OWR model and Section 3.3 presents the results of a horse race between the OWR and the EPIN models.

3.1 Extending the PIN model

Our results in Sections 2.2 and 2.3 show that the PIN model naively identifies information events from turnover. This happens because of two limitations of the PIN model. First, under the PIN model, increases in expected turnover can only come about through the arrival of private information (see Equation 1). Second, the PIN model assumes that order flow is distributed as a mixture of three bivariate Poisson random variables (i.e. the three clusters in Panels A and B of Fig. 3). This assumption is too restrictive to accommodate the high level and volatility of turnover that we observe, especially in the later part of the sample. In this section, we propose an extension of the PIN model to fix the issues with the PIN model.

Before doing so, it is useful to formalize why the model fails in the way that we discuss above. Panel A of Fig. 6 displays a reparameterization of the PIN model in terms of three new parameters. First, the ratio of the intensity of uninformed buyer initiated trades to the intensity of the total number of uninformed trades ($\theta = \epsilon_B/(\epsilon_B + \epsilon_S)$). Second, the ratio of the expected number of informed to uninformed trades on days where there is private information ($\eta = \mu/(\epsilon_B + \epsilon_S)$). Third, the overall intensity of the number of buys plus sells (λ). Specifically, recall that Equation 1 shows that λ is function of the arrival of private information, represented by the indicator $I_{i,t}$ such that on days without private information $\lambda(0) = \epsilon_B + \epsilon_S$ and, on days with private information, $\lambda(1) = \epsilon_B + \epsilon_S + \mu$. The bottom node of Panel A in Fig. 6 shows that, on days without private information, the intensity of buyer initiated trades is $\theta \times \lambda(0)$, while the intensity of seller initiated trades is $(1 - \theta) \times \lambda(0)$. On negative private information days (the central node of Panel A in Fig. 6) the ratio of the intensity of buys to the intensity of total trades drops to $\theta/(1 + \eta)$. Since buy orders are all uninformed and some sell orders are informed, the expected number of buys relative to the expected number of trades is smaller. Finally, on positive information days (the top node of Panel A in Fig. 6) the ratio of sells to the intensity of total trades drops to $(1 - \theta)/(1 + \eta)$. Since sell orders are all uninformed and some buy orders are informed, the expected number of sells relative to the expected number of trades is smaller. Therefore, Panel A of Fig. 6 is a re-parameterization of the PIN model in Fig. 1 using the parameters $\lambda(0)$, η, and θ instead of ϵ_B, ϵ_S, and μ.

Two limitations of the PIN model are immediately clear from the parameterization in Panel A of Fig. 6. First, increases in λ can only come about through the arrival of private information. That is, λ is function of information arrival (I_t). Second, the PIN model does not allow for enough variability in λ to accommodate the high level and volatility of turnover that we observe, especially in the later part of the sample. We resolve the limitations of the PIN model while keeping its information structure with an extension of the PIN model that does two things. First, we draw λ_t independently of the arrival of private information.

Second, we focus on the fraction of trades represented by buys and sells rather than on the absolute amounts of buys and sells following the re-parameterization of the PIN model in Panel A of Fig. 6.

Panel B of Fig. 6 presents the tree structure for the Extended PIN model (EPIN). The EPIN model retains the microstructure intuition of the original PIN model, however, it focuses on the ratios of the expected number of buys and sells to the expected number of trades rather than on the absolute numbers of buys and sells.

Specifically, the EPIN model in Panel B of Fig. 6 draws λ_t from a $Gamma(r, p/(1-p))$ distribution with shape parameter r and scale parameter $p/(1-p)$. The fact that λ_t is drawn from a *Gamma* distribution makes the model particularly tractable since the mixture of the *Poisson* and *Gamma* distributions is the well-known *Negative Binomial* distribution (see Casella and Berger (2002)). In the EPIN model, the number of trades $(B+S)$ is distributed as *Negative Binomial* (see Appendix D for proof), which dramatically simplifies the numerical estimation of the model. In the maximum likelihood estimation the order intensity (λ) parameters r and p can be estimated in a first stage, independently of the remaining information structure parameters which can be estimated in a second stage. $CPIE_{EPIN}$ is calculated in the same way as in the PIN model. Moreover, if we condition down with respect to the data, $CPIE_{EPIN}$ reduces to the model's unconditional probability of information events (α). See Appendix D for a detailed discussion of the model, the associated $EPIN$ measure, the likelihood function, and the $CPIE_{EPIN}$ calculation.

To illustrate how the EPIN model works, we present a stylized example of the EPIN in Fig. 7. Analogous to the PIN model plot in Fig. 3, we plot simulated and real order flow data for Exxon-Mobil during 1993 and 2012, with buys on the horizontal axis and sells on the vertical axis. Panels A and B of Fig. 7 illustrate the central intuition behind the EPIN model. The simulated data comprise three types of days, which create three distinct clusters. Two of the clusters are made up of days characterized by a high proportion of imbalanced trades (large $\frac{|B-S|}{B+S}$), with a large number of sells (buys) and relatively few buys (sells). The third group of days has a low proportion of imbalanced trades–these days have no private information arrival and are clustered around the dashed line in the center of the scatter plots.

The EPIN model implies that days with information events are the ones in which the proportion of imbalanced trades is large. An econometrician using the EPIN model, moving along the dashed line in Panels A and B, would observe that days with above average turnover–days the PIN model classifies as information events–are no longer classified as such, because higher turnover is driven by a large draw of the parameter λ_t under the EPIN model. Instead, the EPIN model identifies private information when moving away from the dashed line–when the proportion of imbalanced trades is high.

Panels C and D plot $CPIE_{EPIN}$ as function of turnover. As opposed to the analogous plot of the PIN model in Fig. 3, Panels C and D do not indicate any relation between turnover and $CPIE_{EPIN}$.[15] Although the EPIN model is not a perfect description of the order flow data, it manages to fix the problem of the PIN model which mechanically identifies private information arrival from turnover.

Table 5 contains summary statistics for the parameter estimates of the EPIN model. Table 5 also contains summary statistics of the cross-sectional sample means and standard deviations of $CPIE_{EPIN}$. We see that the mean $CPIE_{EPIN}$ behaves exactly like α. We also estimate the EPIN model for every stock in our sample in the period $t \in [-312, -60]$ before earnings announcements and opportunistic insider trades. These parameter estimates are used to compute the $CPIE_{EPIN}$ in Section 3.3. The summary statistics of the parameter estimates for the event studies are qualitatively similar to those in Table 5.

3.2 The OWR model

Odders-White and Ready (2008) extend Kyle (1985) by allowing for days with information events and days without information events. Private information arrives before the opening of the trading day with probability α. On days when private information arrives, the model assumes that the information is publically revealed after the close of trade. The OWR model identifies the arrival of private information through order flow imbalance, y_e, the intraday price response to order imbalance, r_d, and through subsequent overnight price changes, r_o.[16]

[15]Internet Appendix D shows the results of regressions of $CPIE_{EPIN}$ on the proportion of imbalanced trades and turnover. These regressions are analogous to those that we performed with the PIN model in Table 3. The results of these regressions indicate that the EPIN model does not conflate turnover with the arrival of private information.

[16]We suppress the t subscript for ease of exposition.

The vector (y_e, r_d, r_o) is assumed to be multivariate normal with mean zero and a covariance matrix that differs between information days and non-information days.[17]

Fig. 8 shows the time line of the model. The intuition behind the OWR model is that the market maker updates prices in response to order flow because the order flow could reflect an information event. However, the subsequent price pattern is different depending on whether there actually was an information event or not. If an information event occurs, the overnight price response reflects a continuation of the market makers' intraday reaction. If no information event occurs, the overnight price response reverses the market makers' initial price reaction. Therefore, an econometrician can make inferences about the probability of an information event in the OWR model because the covariance matrix of the three variables (y_e, r_d, r_o) differs between days when private information arrives and days when only public information is available.[18]

To see how the covariance matrix of (y_e, r_d, r_o) differs between information and non-information days, consider first the covariance of the intraday and overnight returns, $cov(r_o, r_d)$. This covariance is positive for information events, reflecting the fact that the information event is not completely captured in prices during the day and the revelation of the private information means that the overnight return continues the partial intraday price reaction. In contrast, $cov(r_o, r_d)$ is negative in the absence of an information event since the market marker's reaction to the noise trade during the day is reversed when she learns that there was no private signal.

The other moments in the covariance matrix of (y_e, r_d, r_o) are also affected by the arrival of private information. If no information event occurs, then $Var(y_e)$ is composed of only the variances of the uninformed order flow and the noise in the data. However, if an event occurs, $Var(y_e)$ increases because the order flow reflects at least some informed trading. Similarly, $Var(r_d)$ is higher for an information event, because it reflects the market maker's partial reaction to the day's increased order flow. Since the private signal is revealed after

[17]We follow Odders-White and Ready and remove systematic effects from returns to obtain measures of unexpected overnight and intraday returns (r_o and r_d). See Section 1 and Internet Appendix B for a detailed description of how we compute y_e, r_o and r_d.

[18]Unlike the market maker who must update prices before observing the overnight revelation of information, the econometrician in the OWR model can make inferences about the arrival of private information after viewing the overnight price response.

trading closes, $Var(r_o)$ also increases in the wake of an information event, as it reflects the remainder of the market maker's partial reaction to the informed trade component in order flow. Likewise, information events make $cov(y_e, r_d)$ and $cov(y_e, r_o)$ rise. The higher covariance between order flow and intraday returns occurs because, in an information event, both order flow and the intraday return (partially) reflect the impact of informed trading. Along these same lines, because the market maker cannot separate the informed from the uninformed order flow, she is unable to fully adjust the price during the day to reflect the informed trader's private signal. However, since the private signal is publically revealed and fully reflected in prices after the close, $cov(y_e,\ r_o)$ is higher during an information event.

In contrast to the PIN model, the OWR model does not contain a direct analog to the probability of informed trading (PIN). To understand this result, note that the probability of informed trade in the PIN and EPIN models is given by the ratio of the expected number of informed trades to the expected total number of trades on a given day. Since the OWR model employs only the difference between buys and sells, it does not make assumptions about the distribution of number of trades. Thus, the OWR is mute regarding the ratio of the expected number of informed trades to expected number of trades. This may appear to be a limitation of the OWR model, but this is actually an advantage because it allows the OWR model to disentangle variations in turnover from the arrival of informed trading, much like the EPIN model.

Even though the OWR model does not have a measure analogous to the PIN measure, the OWR model admits other useful measures of private information. For instance, the OWR model has a $CPIE_{OWR}$ which reduces to the model's unconditional probability of information events (α) if we condition down with respect to the data. Moreover, Odders-White and Ready (2008) motivate their model as a tool to separate the expected liquidity provider losses due to trading with informed traders into the frequency of private information arrival and the expected magnitude of the private information. Hence, the OWR allows for the construction of private information measures that are based on both dimensions. The PIN and EPIN models, on the other hand, focus only on the frequency of information arrival and are silent with respect to the expected magnitude of the private information. Hence, our comparison of the EPIN and OWR models with $CPIE_{EPIN}$ and $CPIE_{OWR}$ focuses on the dimension of private information that both models have in common, namely the frequency of information arrival. The fact that we are using $CPIE$s to compare the models does not imply that we are taking the position that frequency measures are the only private information metrics that are worthy of consideration.

As with the PIN and EPIN models, we estimate the OWR model numerically via maximum likelihood. Table 6 contains summary statistics for the parameter estimates of the OWR model. Table 6 also contains summary statistics of the cross-sectional sample means and standard deviations of $CPIE_{OWR}$. As in the PIN and EPIN models, we see that the mean $CPIE_{OWR}$ behaves exactly like α in the OWR model. The estimated OWR α parameters are in general higher than those in Odders-White and Ready (2008). This is due to the fact that our definition of y_e is different from that in Odders-White and Ready (2008) (see the discussion in Section 1 above).[19] Fig. 9 plots the time series of the estimated OWR α. In contrast to the PIN α, the OWR α is decreasing over time. This pattern may indicate that private information arrival is less likely later in our sample. While interesting, understanding this pattern is outside the scope of this paper and we leave this investigation for future research. We also estimate the OWR model for each stock i in the period $t \in [-312, -60]$ before earnings announcements and opportunistic insider trades. These parameter estimates are used to compute the $CPIE$s in Sections 3.3.1 and 3.3.2. The summary statistics of the parameter estimates for the event studies are qualitatively similar to those in Table 6. Internet Appendix E has a detailed description of model, its likelihood function, and the $CPIE_{OWR}$ calculation. Appendix E also displays the results of regressions of $CPIE_{OWR}$ similar to those that we perform with $CPIE_{PIN}$ in Section 2.2. These regressions indicate that the OWR model identifies the arrival of private information in a way consistent with its theory.

3.3 A horse race between the EPIN and OWR models

A fundamental problem in the literature related to testing and proposing measures of private information is the lack of cleanly identifiable periods in which private information is present in

[19]In fact, we get α estimates close to those reported in Odders-White and Ready (2008) if we define y_e in the same way that they do.

the market. To address this issue, we use three different methods to analyze the performance of the OWR and EPIN models. In Section 3.3.1 we analyze how $CPIE_{OWR}$ and $CPIE_{EPIN}$ vary around earnings announcements. The assumption underlying this test is that private information arrival is more likely before than after the announcement. In Section 3.3.2 we analyze how $CPIE_{OWR}$ and $CPIE_{EPIN}$ vary around insider trading events. In Section 3.3.3 we analyze how $CPIE_{OWR}$ and $CPIE_{EPIN}$ are related to return autocorrelations.

Each of these three methods has its own unique limitations. For instance, it is possible that, for some reason, private information is more prevalent after important announcements than before. Other critiques could be levied against the other two methods. However, if all of these methods point to the same conclusions, it seems unlikely that our overall interpretation would be biased due to the limitations of any specific method.

3.3.1 Information event probabilities under the EPIN and OWR models

Panel A of Fig. 10 illustrates the average $CPIE_{EPIN}$ in event time for our sample of earnings announcements. In contrast to the PIN model, the probability of an information event decreases from around 51% 20 days before the announcement and drops on the announcement date to around 46%. This pattern is not consistent with informed traders acting on private information before the announcement. Panel B of Fig. 10 illustrates the average $CPIE_{OWR}$ in event time for our sample of earnings announcements. Similar to the PIN model, the probability of an information event increases from around 40% 20 days before the announcement and peaks on the announcement date at around 45%. Panel B indicates that the $CPIE_{OWR}$ is far outside of two standard deviations from its mean (estimated between $t \in [-40, -21]$) on the announcement date $t = 0$. This pattern is consistent with the timing of the OWR model where informed traders act on private information during the day before the public announcement which occurs overnight ($t \in [0, 1)$). Unlike the PIN model, the $CPIE_{OWR}$ drops back to its pre-event mean within a few days after the announcement. This is consistent with the intuition that there is more scope for informed trading before the announcement than after.

What causes the EPIN results to be so different from the PIN results above? Fig. 11 sheds light on this question. Panel A of Fig. 11, shows the actual $CPIE_{EPIN}$ along with predicted values from a regression of $CPIE_{EPIN}$ on the proportion of imbalanced trades $\left(\frac{|B-S|}{B+S}\right)$ and its square. The results indicate that $CPIE_{EPIN}$ drops because the imbalance is small relative to the absolute amount of trade on the announcement day. This is consistent with the results in Easley, Engle, O'Hara, and Wu (2008), who show that in their sample of 834 announcements that the average proportion of imbalanced trades decreases on earnings announcement days. The PIN model interprets the increase in turnover as indicative of the arrival of private information, but the EPIN model, on the other hand, uses the information in the proportion of imbalanced trades to draw the opposite conclusion. Panel B provides support for this notion by showing that the $CPIE_{EPIN}$ does not respond to increases in turnover. Panel B shows the predicted $CPIE_{EPIN}$ based on a regression of $CPIE_{EPIN}$ on $\frac{|B-S|}{B+S}$ and its square. The results indicate that, consistent with the motivation for the extended model, $CPIE_{EPIN}$ responds to the proportion of imbalanced trades and not turnover.

As we saw in Section 3.2, the OWR model identifies private information from the covariance matrix of the three variables in the model $(y_{e,i,t}, r_{o,i,t}, r_{d,i,t})$. Therefore, to analyze how the OWR model identifies private information around earnings announcements, we decompose $CPIE_{OWR}$ on to the squared and interaction terms of $(y_{e,i,t}, r_{o,i,t}, r_{d,i,t})$. Panels A–F of Fig. 12 show that the majority of the variation in measured private information ($CPIE_{OWR}$) comes from intraday returns squared (Panel B) and the interaction between the intraday and overnight returns (Panel F). Order imbalance squared (Panel A) provides no explanatory power, although the interaction between the order imbalance and returns (Panels D and E) does have some impact.

Our results suggest that order flow, however well modeled, is insufficient to be the sole source of inferences about private information arrival. Under the assumption that there is more informed trade before rather than after earnings announcements, our findings suggest that the OWR model identifies private information in a sensible way while the EPIN does not. Even though the magnitude of the increase in $CPIE_{OWR}$ around the event date may be considered small, $CPIE_{OWR}$ increases before the event day while $CPIE_{EPIN}$ counter-intuitively decreases. Since both models use order flow to identify private information, the marked difference in the results highlights the importance of including the price response mechanism. The use of returns, particularly intraday returns, allows the OWR model to reach a different and more economically sensible conclusion. Moreover, the fact that order imbalance alone explains very little of the variation in $CPIE_{OWR}$ around earnings announcements also emphasizes the relatively low contribution of order flow relative to returns in identifying private information. Our results therefore provide empirical support for the proposition in Back, Crotty, and Li (2014) and in Kim and Stoll (2014) that researchers cannot use order flow alone to successfully identify periods of informed trade.

3.3.2 $CPIE_{EPIN}$ and $CPIE_{OWR}$ around insider trading

In this section we investigate whether the OWR and EPIN models are capable of identifying opportunistic insider trades using the insider trade classification scheme developed in Cohen, Malloy, and Pomorski (2012).[20] Cohen, Malloy, and Pomorski (2012) show that a long-short portfolio that exploits the trades of opportunistic traders (opportunistic buys minus opportunistic sells) earns value-weighted abnormal returns of 82 basis points per month (9.8 percent annualized, t-statistic=2.15). They also show that the trades of opportunistic insiders show significant predictive power for future news about the firm, and that the fraction of traders who are opportunistic in a given month is negatively related to the number of recent news releases by the SEC regarding illegal insider trading cases. Their results are all consistent with opportunistic insider trades, as opposed to routine insider trades, being based on private information. Opportunistic insider trades therefore, provide a convenient laboratory to examine the models' ability to detect the arrival of actionable private information.

Panel A (B) of Fig. 13 presents the average $CPIE_{EPIN}$ ($CPIE_{OWR}$) in event time for our sample of opportunistic insider trades. There is no clear pattern in the $CPIE_{EPIN}$ indicating the arrival of private information before opportunistic insider trades, though there is an increase in $CPIE_{EPIN}$ on the day of opportunistic insider trades.

In contrast, Panel B shows that the $CPIE_{OWR}$ identifies the arrival of private information in the days leading up to an opportunistic insider trade. Beginning at $t = -4$, the $CPIE_{OWR}$ is more than two standard deviations higher than the mean estimated between $t \in [-40, 21]$. However, $CPIE_{OWR}$ begins to drift strongly upward and very nearly crosses the two standard deviation bound as early as day $t = -10$. Strikingly, at $t = 1$, immediately after the trade,

[20] See Section 1 for a further discussion of the classification of insider trades as opportunistic.

$CPIE_{OWR}$ drops precipitously back to average levels. We interpret this as strong evidence that the OWR model's use of both order flow and returns is successful in uncovering informed trade.

Taken together, the insider trading event study evidence further supports the claim that order flows alone may be insufficient to identify private information. $CPIE_{EPIN}$, which varies based only on changes in order imbalances, is unable to clearly detect the imminent arrival of insider trades. $CPIE_{OWR}$, on the other hand, is able to predict insider trading based on small variations in intraday and overnight returns.

3.3.3 Are $CPIE_{EPIN}$ and $CPIE_{OWR}$ related to return continuation?

The market microstructure literature has long held that price changes related to informed trades should not be subsequently reversed while non-information related price changes (e.g. dealer inventory control, price pressure, price discreteness etc.) are transient (e.g. Hasbrouck (1988, 1991a,b)). In this section, we investigate whether $CPIE_{EPIN}$ and $CPIE_{OWR}$ are associated with subsequent return reversals. In particular, we examine the relation between $CPIE$s and return autocorrelations. The intuition is that if a model's $CPIE$ on day t actually reflects a high probability of informed trade then we expect that the return on day t will be continued over the subsequent day as the information gradually becomes public and gets fully impounded in prices. To capture this idea we model return autocorrelations as linear functions of $CPIE$. Specifically, we consider the following regressions: $r_{i,t+1} = \alpha + \beta_{OWR,1} r_{i,t} + \beta_{OWR,2} CPIE_{OWR,t} + \beta_{OWR,3}(r_{i,t} \times CPIE_{OWR,t}) + \varepsilon_{i,t+1}$, and $r_{i,t+1} = \alpha + \beta_{EPIN,1} r_{i,t} + \beta_{EPIN,2} CPIE_{EPIN,t} + \beta_{EPIN,3}(r_{i,t} \times CPIE_{EPIN,t}) + \varepsilon_{i,t+1}$.

In the above regressions, $r_{i,t}$ is the open-to-open, risk adjusted return $(r_{i,d,t} + r_{i,o,t})$ on day t. Thus, there is no overlap between the intraday and overnight returns that are used to compute $CPIE_{OWR,i,t}$ on day t and the return on day $t+1$. The coefficients $\beta_{OWR,2}$ and $\beta_{EPIN,2}$ reflect the impact of the model's $CPIE$ on the correlation between the return on day t and the return the next trading day. We estimate the regressions above using a panel regression approach including firm and year fixed effects with standard errors clustered by firm and year. Table 7 reports the coefficient estimates and t-statistics for these regressions. The results in Table 7 show that the estimates for both $\beta_{OWR,2}$ and $\beta_{EPIN,2}$ are positive and significant, indicating that both $CPIE_{EPIN}$ and $CPIE_{OWR}$ are associated with future return continuation. To see this note that both regressions show a tendency of daily returns to reverse because the coefficients on lagged returns in both regressions are negative. However, a one standard deviation shock to $CPIE_{OWR}$ is associated with a 0.02 (0.08×0.25) decline in the subsequent reversal, while a one standard deviation shock to $CPIE_{EPIN}$ is associated with only a 0.003 (0.006×0.49) drop in the subsequent reversal. Thus, while the point estimates for both the OWR and EPIN models suggest that $CPIE_{EPIN}$ and $CPIE_{OWR}$ capture information that has a persistent impact on prices, the effect is ten times stronger with the OWR $CPIE$. We view this as further evidence that including the price response mechanism allows researchers to make stronger inferences about private information arrival.

4 Conclusion

The PIN measure, developed in the seminal work of Easley and O'Hara (1987), Easley, Kiefer, O'Hara, and Paperman (1996), and Easley, Kiefer, and O'Hara (1997), is arguably the most widely used measure of information asymmetry in the accounting, corporate finance and asset pricing literature today. Recent work however suggests that PIN fails to capture private information (e.g. Aktas, de Bodt, Declerck, and Van Oppens (2007), Benos and Jochec (2007), and Collin-Dufresne and Fos (2014a)). This paper analyzes why the model might incorrectly identify informed trade.

Our findings indicate that the PIN model fits the data so poorly that it mechanically groups *all* sources of variation in turnover (e.g. disagreement, calendar effects, portfolio rebalancing, taxation, etc.) under the umbrella of private information arrival. This is at odds with a vast literature that suggests turnover varies for many reasons unrelated to the arrival of private information. This failure of the PIN model is particularly strong after the increase in turnover in the early 2000s. In fact, after 2002 for the median stock in our sample, the PIN model is essentially equivalent to a naïve model that assigns a probability of one to the arrival of private information on any day where turnover is above average and zero probability to the arrival of private information on any other day. These findings suggest some important insights for future research that tests, constructs, or uses proxies for informed trade.

Our results suggest that event study based tests of private information proxies (e.g. Easley, Engle, O'Hara, and Wu (2008) and Brennan, Huh, and Subrahmanyam (2015)) can be misleading if one fails to account for the fact that patterns in private information measures may simply reflect event-related patterns in turnover that have nothing to do with private information arrival. For instance, Brennan, Huh, and Subrahmanyam (2015) interpret the fact that their $CPIE_{PIN}$ measures are higher after earnings announcements than before as evidence of informed trading. However, we show that $CPIE_{PIN}$ is mechanically related to turnover. This suggests that the findings in Brennan, Huh, and Subrahmanyam (2015) can simply be attributed to the fact that turnover is typically much higher after earnings announcements.

Our findings also suggest that future research aimed at building measures of informed trade should focus on the price response mechanism in addition to net order because order flow, however well modeled, appears insufficient to identify private information. Specifically, we use three different methods to compare the OWR model, which infers the arrival of private information from returns and order flow, with an extension of the PIN model (the EPIN model), which is solely based on order flow but corrects the PIN model's mechanical association of private information arrival with variation in turnover. The OWR model performs better than the EPIN model in all three tests. First, the EPIN model actually predicts a decrease in private information arrival before earnings announcements while the OWR model captures a pattern of increasing private information arrival prior to the announcement and a marked decrease after the announcement. Second, $CPIE_{OWR}$ predicts periods of opportunistic insider trading and decreases dramatically immediately following the insider trades, while $CPIE_{EPIN}$ displays no such clear pattern around these events. Lastly, the relation between $CPIE_{OWR}$ and future return continuation is ten times larger than that of the $CPIE_{EPIN}$.

Our findings also suggest that future research in corporate finance, accounting, or asset pricing that uses information asymmetry measures should consider using proxies for private information based on the OWR model, for instance $CPIE_{OWR}$ or its α, instead of using proxies based on the PIN model (e.g. PIN).

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Table 1: **Summary Statistics.** This table summarizes the full sample and event day (t=0) returns, order imbalance, and number of buys and sells. We compute intraday and overnight returns as well as daily buys and sells for stocks between 1993 and 2012 using data from the NYSE TAQ database. Following Odders-White and Ready (2008), we compute the intraday return, r_d, at time t as the volume-weighted average price at t (VWAP) minus the opening quote midpoint at t plus dividends at time t, all divided by the opening quote midpoint at time t. We compute the overnight return, r_o, at t as the opening quote midpoint at $t+1$ minus the VWAP at t, all divided by the opening quote midpoint at t. We compute y_e as the daily total volume of buys minus total volume of sells, divided by the total volume. For the PIN and EPIN models, we use the daily total number of buys and sells. Our sample of earnings announcements includes all CRSP/COMPUSTAT firms listed in NYSE between 1995–2009 for which we have exact timestamps collected from press releases in Factiva which fall within a [-1,0] window relative to COMPUSTAT earnings announcement dates. Opportunistic insider trades are defined as in Cohen, Malloy, and Pomorski (2011).

(a) Full Sample

	N	Mean	Std	Q1	Median	Q3
y_e	5,286,191	2.766%	31.259%	-10.433%	3.282%	18.996%
r_d	5,286,191	-0.004%	1.500%	-0.707%	-0.024%	0.680%
r_o	5,286,191	0.003%	1.297%	-0.566%	-0.024%	0.525%
# Buys	5,286,191	1,876	6,917	37	220	1,128
# Sells	5,286,191	1,843	6,894	36	194	1,033

(b) Earnings Announcements

	N	Mean	Std	Q1	Median	Q3
y_e	21,979	5.099%	22.122%	-4.787%	4.373%	16.400%
r_d	21,979	0.002%	2.424%	-1.252%	-0.004%	1.271%
r_o	21,979	0.075%	2.313%	-1.042%	0.013%	1.153%
# Buys	21,979	4,572	13,491	223	956	3,421
# Sells	21,979	4,465	13,546	191	831	3,165

(c) Opportunistic Insider Trades

	N	Mean	Std	Q1	Median	Q3
y_e	32,676	4.980%	20.425%	-5.106%	3.874%	15.353%
r_d	32,676	0.151%	1.566%	-0.632%	0.086%	0.865%
r_o	32,676	0.056%	1.247%	-0.467%	0.020%	0.528%
# Buys	32,676	3,852	10,645	354	1,129	3,478
# Sells	32,676	3,787	10,554	300	996	3,303

Table 2: **PIN Parameter Estimates.** This table summarizes parameter estimates of the PIN model for 21,206 `PERMNO`-Year samples from 1993 to 2012. α represents the average unconditional probability of an information event at the daily level. δ represents the probability of good news, and $1-\delta$ represents the probability of bad news. ϵ_B and ϵ_S represent the expected number of daily buys and sells given no private information. μ represents the expected additional order flows given an information event. $\overline{CPIE}$ and Std($CPIE$) are the `PERMNO`-Year mean and standard deviation of $CPIE_{PIN}$.

	N	Mean	Std	Q1	Median	Q3
α	21,206	0.372	0.122	0.291	0.375	0.445
δ	21,206	0.607	0.209	0.484	0.625	0.762
ϵ_B	21,206	1,625	5,388	33	193	1,039
ϵ_S	21,206	1,596	5,369	35	186	956
μ	21,206	312	593	43	160	314
$\overline{CPIE}$	21,206	0.382	0.135	0.293	0.379	0.449
Std($CPIE$)	21,206	0.451	0.052	0.427	0.470	0.490

Table 3: **PIN Model Regressions.** This table reports real and simulated regressions of the $CPIE_{PIN}$ on absolute order imbalance ($|B-S|$), and order imbalance squared ($|B-S|^2$). In Panel A, we simulate 1,000 instances of the PIN model for each `PERMNO`-Year in our sample (1993–2012) and report mean standardized estimates for the median stock, along with 5%, 50%, and 95% values of the R^2 ($R^2_{inc.}$) values. We compute the incremental $R^2_{inc.}$ as the R^2 attributed to $turn$ and $turn^2$ in an extended regression model. In Panel B, we report standardized estimates for the median stock using real data, along with the median R^2 and $R^2_{inc.}$ values, and tests of the null hypothesis that the observed relation between $CPIE_{PIN}$ and $turn$ is consistent with the PIN model. The p-value of is the mean probability under the null of observing an $R^2_{inc.}$ at least as large as what is observed in the real data. The % Rej. is the fraction of stocks for which we reject the null hypothesis at the 5% level.

(a) Simulated Data

	β		t		R^2			$R^2_{inc.}$		
	$\|B-S\|$	$\|B-S\|^2$	$\|B-S\|$	$\|B-S\|^2$	5%	50%	95%	5%	50%	95%
1993	0.437	-0.079	(10.31)	(-1.80)	71.13%	76.09%	80.38%	7.17%	10.57%	15.25%
1994	0.422	-0.072	(9.63)	(-1.67)	67.49%	73.26%	78.11%	9.39%	13.47%	18.55%
1995	0.410	-0.058	(9.68)	(-1.36)	70.32%	75.39%	79.85%	7.64%	11.39%	16.02%
1996	0.432	-0.085	(9.89)	(-1.90)	69.02%	74.28%	78.87%	8.32%	12.17%	16.97%
1997	0.450	-0.089	(10.30)	(-1.98)	71.99%	76.93%	81.12%	7.36%	10.76%	14.79%
1998	0.482	-0.104	(10.79)	(-2.36)	74.32%	78.71%	82.46%	6.65%	9.53%	13.30%
1999	0.484	-0.112	(11.03)	(-2.47)	75.62%	79.96%	83.46%	6.49%	9.36%	12.92%
2000	0.529	-0.137	(11.88)	(-3.00)	79.78%	83.36%	86.15%	4.98%	7.47%	10.45%
2001	0.638	-0.217	(13.97)	(-4.61)	83.34%	86.13%	88.57%	4.17%	6.00%	8.35%
2002	0.695	-0.260	(14.11)	(-5.30)	82.61%	85.53%	88.06%	4.83%	6.92%	9.54%
2003	0.665	-0.244	(12.38)	(-4.52)	78.88%	82.36%	85.36%	7.90%	10.56%	13.79%
2004	0.650	-0.223	(11.49)	(-4.16)	77.84%	81.38%	84.59%	8.92%	11.67%	15.03%
2005	0.658	-0.220	(12.59)	(-4.46)	80.47%	83.59%	86.45%	7.69%	10.09%	12.95%
2006	0.650	-0.221	(11.96)	(-4.35)	80.31%	83.36%	86.18%	7.76%	10.29%	13.50%
2007	0.632	-0.222	(9.40)	(-4.07)	79.72%	83.35%	86.15%	8.53%	10.93%	14.05%
2008	0.666	-0.235	(12.29)	(-4.83)	82.44%	85.25%	88.00%	6.83%	9.15%	11.78%
2009	0.709	-0.269	(14.37)	(-5.70)	84.29%	86.87%	89.20%	6.22%	8.28%	10.57%
2010	0.704	-0.261	(14.60)	(-5.68)	84.99%	87.41%	89.64%	5.66%	7.55%	9.89%
2011	0.671	-0.234	(14.13)	(-5.21)	85.91%	88.25%	90.21%	5.34%	7.28%	9.39%
2012	0.693	-0.251	(14.92)	(-5.62)	85.68%	87.98%	90.34%	5.22%	7.22%	9.50%

Table 3: **PIN Model Regressions.** Continued.

(b) Real Data

	β		t		R^2	$R^2_{inc.}$		
	$\|B-S\|$	$\|B-S\|^2$	$\|B-S\|$	$\|B-S\|^2$	50%	50%	p-value	% Rej.
1993	0.300	-0.073	(5.98)	(-1.43)	35.76%	36.20%	2.57%	94.07%
1994	0.264	-0.047	(5.28)	(-0.92)	32.82%	40.02%	3.36%	92.17%
1995	0.280	-0.061	(5.77)	(-1.29)	34.20%	36.97%	5.05%	89.29%
1996	0.277	-0.065	(5.69)	(-1.28)	30.92%	38.97%	3.85%	92.30%
1997	0.283	-0.073	(5.67)	(-1.36)	30.80%	38.86%	3.54%	92.99%
1998	0.274	-0.059	(5.26)	(-1.09)	30.12%	39.58%	3.54%	93.67%
1999	0.280	-0.059	(5.21)	(-1.08)	29.05%	39.46%	3.29%	94.29%
2000	0.300	-0.079	(5.48)	(-1.39)	29.99%	39.08%	2.59%	95.63%
2001	0.339	-0.111	(5.67)	(-1.87)	29.44%	39.39%	3.53%	94.76%
2002	0.279	-0.058	(4.09)	(-0.85)	23.05%	44.28%	5.59%	91.48%
2003	0.247	-0.032	(3.57)	(-0.47)	21.97%	41.86%	9.55%	84.87%
2004	0.211	-0.005	(3.14)	(-0.08)	19.55%	45.22%	8.78%	86.21%
2005	0.254	-0.053	(3.81)	(-0.81)	19.42%	46.29%	9.21%	85.47%
2006	0.251	-0.066	(3.80)	(-0.96)	16.95%	48.44%	10.83%	85.30%
2007	0.271	-0.104	(4.01)	(-1.57)	14.30%	50.32%	14.04%	82.00%
2008	0.268	-0.111	(4.00)	(-1.66)	13.78%	50.97%	11.49%	86.08%
2009	0.280	-0.117	(4.15)	(-1.74)	14.59%	49.91%	10.08%	87.58%
2010	0.291	-0.124	(4.39)	(-1.82)	15.96%	47.64%	10.62%	87.45%
2011	0.295	-0.131	(4.56)	(-2.03)	15.94%	46.60%	11.14%	86.90%
2012	0.319	-0.145	(4.96)	(-2.23)	17.56%	45.61%	13.31%	85.12%

Table 4: **PIN Regressions Around Earnings Announcements.** This table reports regression results for $CPIE_{PIN}$ around Earnings Announcements. For each announcing firm in our sample we run regressions of $CPIE_{PIN}$ on absolute order imbalance ($|B-S|$) and absolute order imbalance squared ($|B-S|^2$) from $[-20,+20]$ and report median estimates across all the events. We compute the incremental $R^2_{\text{inc.}}$ as the increase in R^2 attributed to *turn* and $turn^2$ in an extended regression model. We report standardized coefficients.

β		t		R^2	$R^2_{\text{inc.}}$
$\lvert B-S\rvert$	$\lvert B-S\rvert^2$	$\lvert B-S\rvert$	$\lvert B-S\rvert^2$	50%	50%
0.143	-0.032	(1.07)	(-0.35)	15.42%	44.44%

Table 5: **EPIN Parameter Estimates.** This table summarizes parameter estimates of the EPIN model for 21,206 `PERMNO`-Year samples from 1993 to 2012. α represents the average unconditional probability of an information event at the daily level. δ represents the probability of good news, and $1-\delta$ represents the probability of bad news. The total number of trades in any given day (t) is drawn from a Poisson distribution with intensity λ_t, where λ_t is draw from a Gamma distribution with shape parameter r and scale parameter $p/(1-p)$. The number of buys on a day with no private information is draw from a Poisson distribution with intensity $\theta \times \lambda_t$. On days with negative news, the number of buys is drawn from a Poisson with intensity $\theta/(1+\eta) \times \lambda_t$. $\overline{CPIE}$ and Std($CPIE$) are the `PERMNO`-Year mean and standard deviation of $CPIE_{EPIN}$.

	N	Mean	Std	Q1	Median	Q3
α	21,206	0.493	0.088	0.448	0.498	0.543
δ	21,206	0.495	0.184	0.372	0.492	0.616
r	21,206	7.210	4.724	4.056	5.976	8.960
p	21,206	0.948	0.080	0.932	0.984	0.997
θ	21,206	0.515	0.049	0.493	0.514	0.546
η	21,206	0.316	0.242	0.152	0.240	0.413
$\overline{CPIE}$	21,206	0.494	0.087	0.449	0.499	0.543
Std($CPIE$)	21,206	0.414	0.082	0.367	0.445	0.478

Table 6: **OWR Parameter Estimates.** This table summarizes parameter estimates of the OWR model for 21,206 `PERMNO`-Year samples from 1993 to 2012. α represents the average unconditional probability of an information event at the daily level. σ_u represents the standard deviation of the order imbalance due to uninformed traders, which is observed with normally distributed noise with variance σ_z^2. σ_i represents the standard deviation of the informed trader's private signal. σ_{pd} and σ_{po} represent the standard deviation of intraday and overnight returns, respectively. $\overline{CPIE}$ and Std($CPIE$) are the `PERMNO`-Year mean and standard deviation of $CPIE_{OWR}$.

	N	Mean	Std	Q1	Median	Q3
α	21,206	0.437	0.257	0.214	0.436	0.639
σ_u	21,206	0.075	0.068	0.022	0.062	0.109
σ_z	21,206	0.239	0.143	0.137	0.221	0.332
σ_i	21,206	0.030	0.286	0.013	0.021	0.027
σ_{pd}	21,206	0.010	0.005	0.006	0.009	0.012
σ_{po}	21,206	0.006	0.004	0.004	0.006	0.008
$\overline{CPIE}$	21,206	0.451	0.258	0.227	0.455	0.656
Std($CPIE$)	21,206	0.137	0.047	0.109	0.142	0.171

Table 7: **Return Reversals.** This table reports regressions of the daily return at time $t+1$ on the return, $CPIE$ ($CPIE_{EPIN}$ or $CPIE_{OWR}$), and the interaction at time t. Returns are measured from open to open and they are computed as the sum of the intraday (r_d) and overnight returns (r_o). We include stock and year fixed effects and cluster standard errors by stock and year. * indicates statistical significance at the 10% level, ** at the 5%, and *** at the 1% level.

	r_{t+1}	
	OWR	EPIN
r_t	-8.883***	-6.955***
	(-6.88)	(-6.91)
CPIE$_t$	0.0136***	
	(4.36)	
CPIE$_t$ × r_t	2.417***	
	(4.16)	
CPIE$_t$		0.00704***
		(4.03)
CPIE$_t$ × r_t		0.271**
		(2.58)
R^2(%)	0.61	0.54
Obs.	5,284,078	5,284,078

Figure 1: **PIN Tree.** For a given trading day, private information arrives with probability α. When there is no private information, buys and sells are Poisson with intensity ϵ_B and ϵ_S. Private information is good news with probability δ. The expected number of buys (sells) increases by μ in case of good (bad) news.

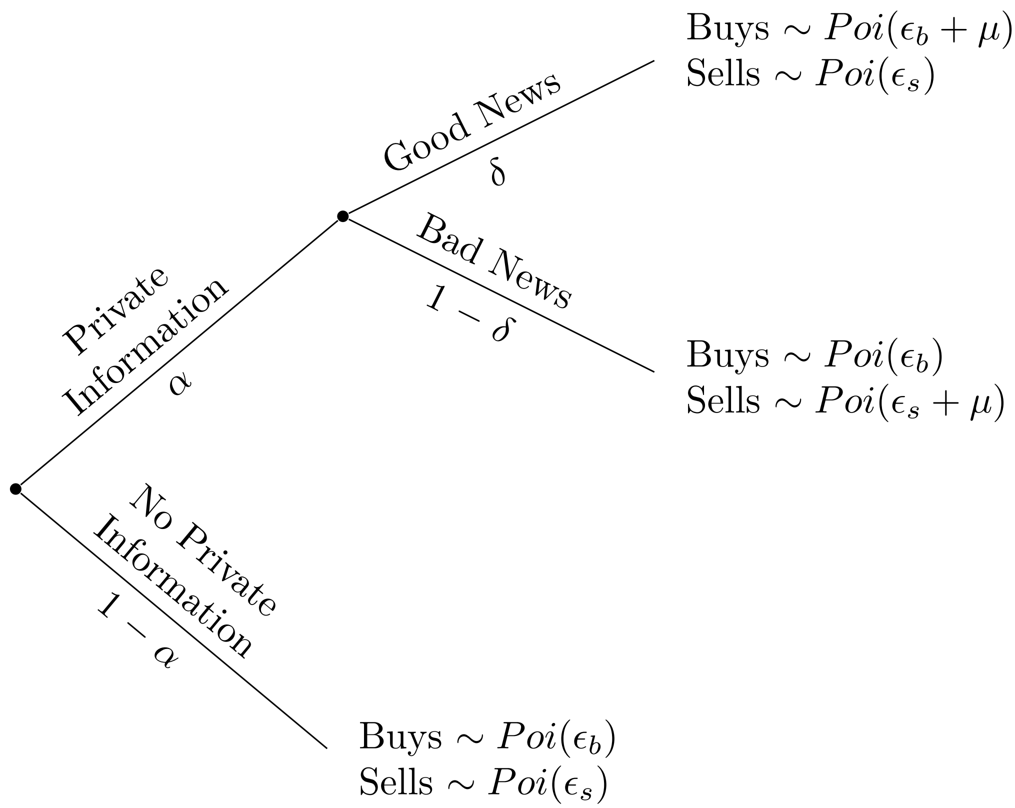
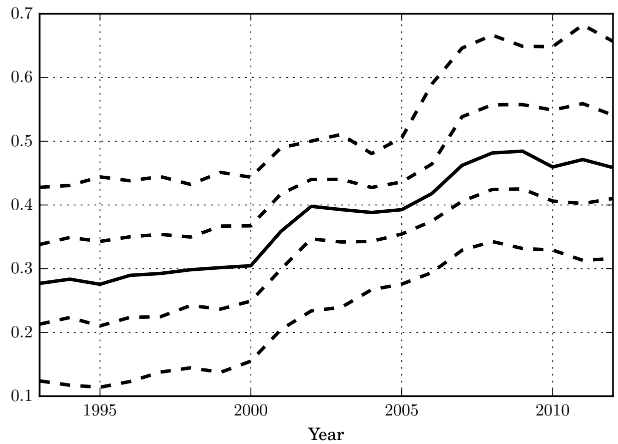
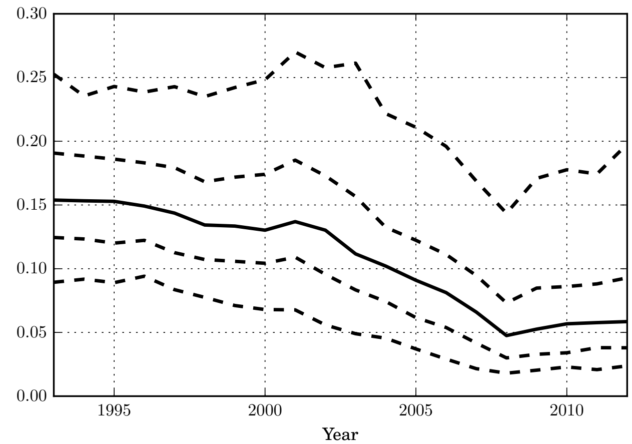


Figure 2: PIN **Parameters.** This figure shows the distribution of yearly α, PIN, and $\mu, \epsilon_B, \epsilon_S$ parameter estimates for the PIN model. The solid black line represents the median value, and the dotted lines represent the 5, 25, 75, and 95 percentiles.

(a) PIN α



(b) PIN



(c) PIN Parameters

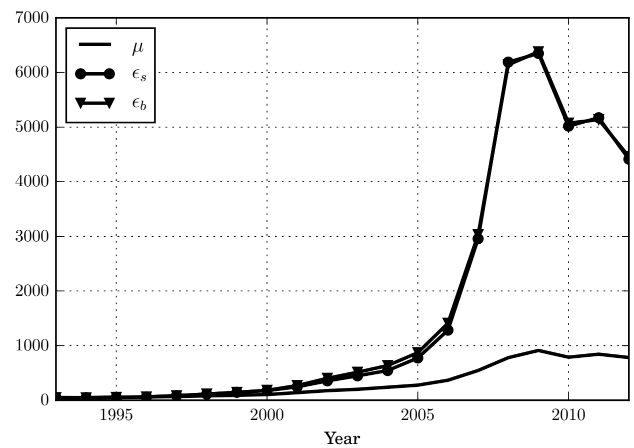
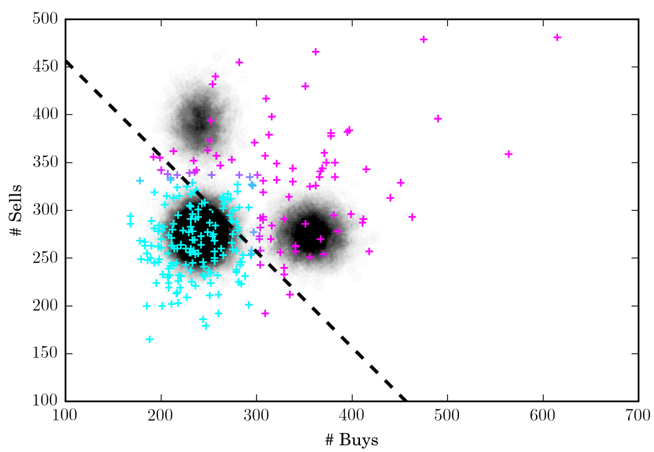
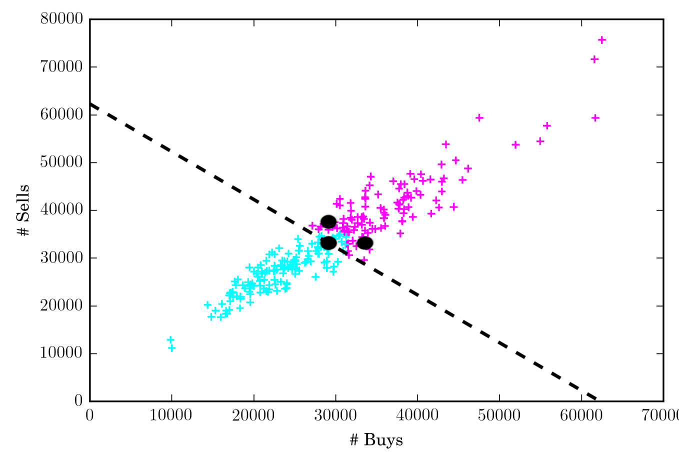


Figure 3: **XOM EO.** This figure compares the real and simulated data for XOM in 1993 and 2012 using the PIN model. In Panels A and B, the real data are marked as +. The real data are shaded according to the $CPIE_{PIN}$, with darker markers (+ magenta) representing high and lighter markers (+ cyan) low $CPIE$s. High (low) probability states in the simulated data appear as a dark (light) "cloud" of points. The PIN model has three states: no news, good news, and bad news. All the observations below (above) the dashed lines in Panels A and B have turnover below (above) the annual mean of daily turnover. Panels C and D plot the CPIEs for the real data as a function of turnover along with a dashed line indicating the mean turnover.

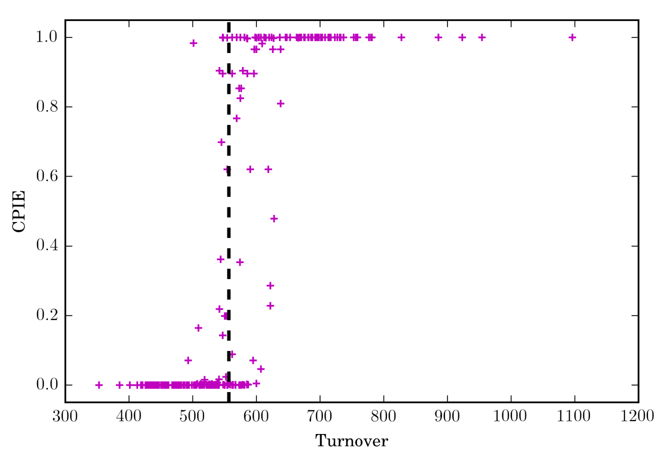
(a) XOM 1993



(b) XOM 2012



(c) XOM 1993



(d) XOM 2012

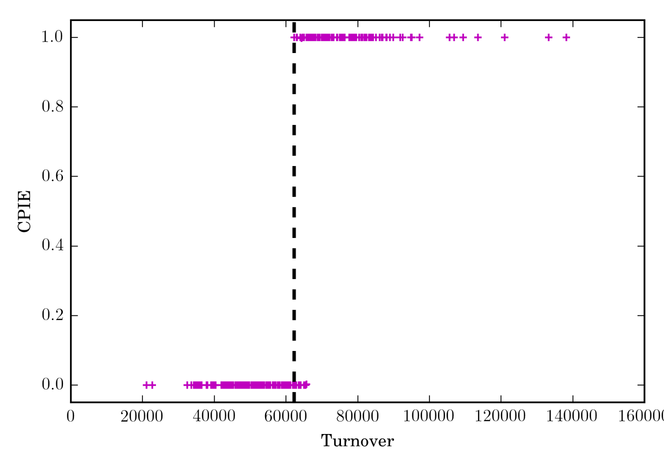
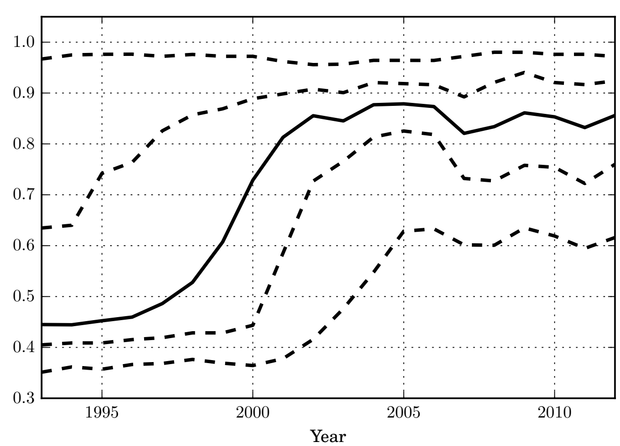


Figure 4: **Breakdown of the PIN Model.** Panel A shows the distribution of the percent of trading days in a year in which the PIN model identifies private information essentially in the same way as the naive identification scheme. That is, Panel A plots the percentage of days where the $|CPIE_{PIN} - CPIE_{Naive}| < 10^{-10}$. $CPIE_{Naive}$ is one for a given stock-day if turnover is higher than the annual mean of daily turnover, and is zero otherwise. Panel B shows the distribution of the percent of days where the likelihood, given the model parameters and observed order flow data is less than 10^{-10} — days, according to the model, with near-zero probability of occurring. The solid black line represents the median stock, and the dashed lines represent the 5, 25, 75, and 95 percentiles.

(a) Days with $CPIE_{PIN} \approx CPIE_{Naive}$



(b) Days with Near-Zero Probability

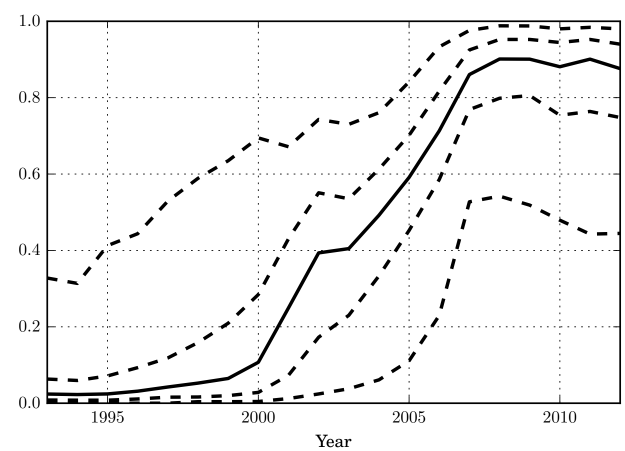
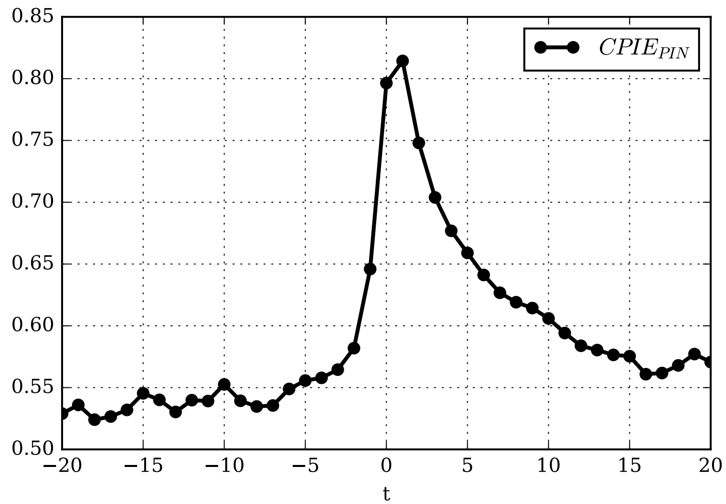
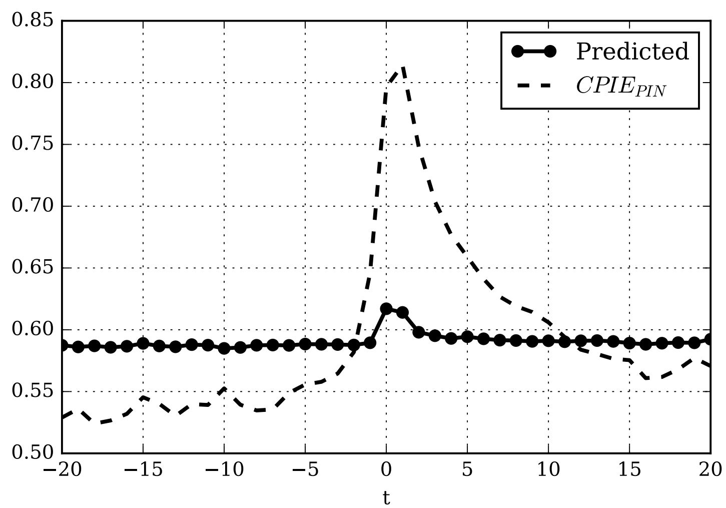


Figure 5: **Earnings Announcements - PIN.** Panel A shows the average $CPIE_{PIN}$ for the PIN model in event time surrounding earnings announcements. Panels B and C compare the average $CPIE_{PIN}$ with the $CPIE_{PIN}$ predicted with either the absolute order imbalance ($|B-S|$) or turnover ($turn$), respectively. To obtain the predictions, we run regressions of daily $CPIE_{PIN}$ on $|B-S|$ or $turn$, and their respective squared terms.

(a) $CPIE_{PIN}$



(b) Prediction using $|B-S|$ and $|B-S|^2$



(c) Prediction using $turn$ and $turn^2$

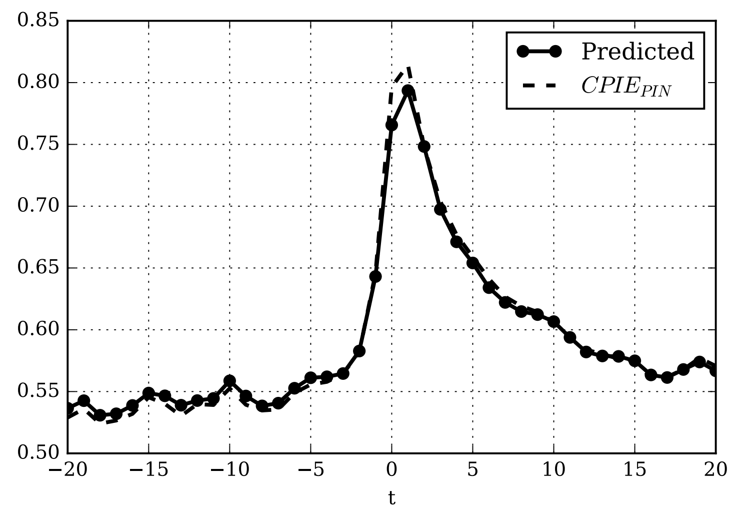
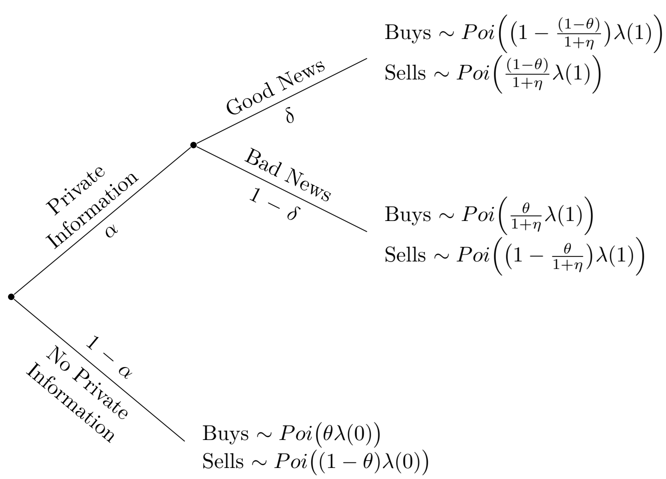


Figure 6: **EPIN Tree.** Panel A presents a re-parameterization of the PIN model in terms of ratio of the intensity of uninformed buyer initiated trades to the intensity of the total number of uninformed trades ($\theta = \epsilon_B/(\epsilon_B + \epsilon_S)$), the ratio of the expected number of informed to uninformed trades on days where there is private information ($\eta = \mu/(\epsilon_B + \epsilon_S)$), and the overall intensity of the number of buys plus sells as a function of the arrival of private information ($\lambda(I_{i,t})$). Panel B presents the EPIN model. The EPIN model extends the PIN model by allowing the intensity of the number of trades on a given day t (λ_t) to be drawn from a Gamma distribution with location and scale parameters r and $p/(1-p)$, respectively. The information structure remains the same as the one in the PIN model. For a given trading day, private information arrives with probability α. When there is no private information, the number of buys (sells) is distributed as a Poisson with intensity $\theta \times \lambda_t$ $\big((1-\theta)\times\lambda_t\big)$. Private information is good (bad) news with probability δ $(1-\delta)$. When there is good news, the number of sells (buys) is Poisson with intensity $\frac{(1-\theta)}{1+\eta}\lambda_t$ $\big((1-\frac{(1-\theta)}{1+\eta})\lambda_t\big)$. When there is bad news, the number of buys (sells) is Poisson with intensity $\frac{\theta}{1+\eta}\lambda_t$ $\big((1-\frac{\theta}{1+\eta})\lambda_t\big)$.

(a) PIN Re-parameterization



(b) EPIN Tree

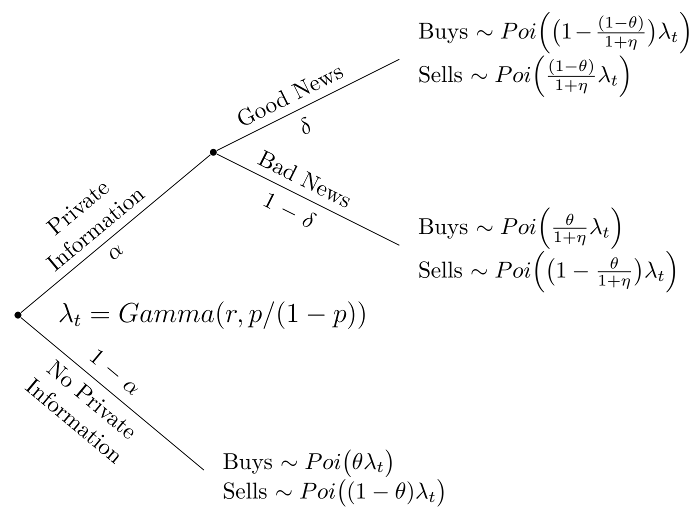
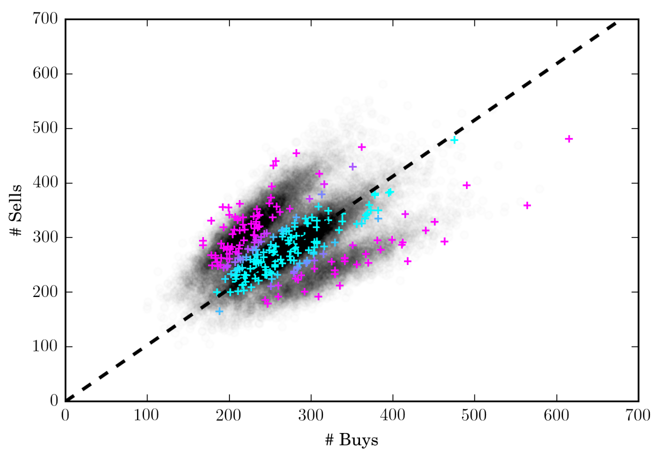
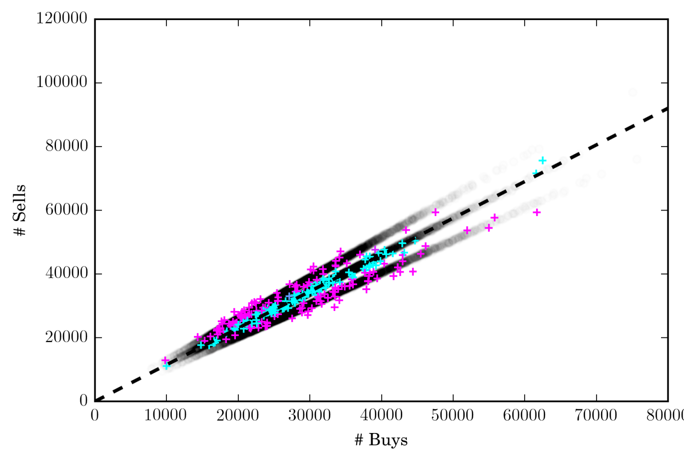


Figure 7: **XOM EPIN.** This figure compares the real and simulated data for XOM in 1993 using the EPIN model. In Panels A and B, the real data are marked as +. The real data are shaded according to the $CPIE_{EPIN}$, with darker markers (+ magenta) representing high and lighter markers (+ cyan) low $CPIE$s. The simulated data points are represented by transparent dots, such that high probability states appear as a dense, dark "cloud" of points, and low probability states appear as a light "cloud" of points. The EPIN model has three states: no news, good news, and bad news. Panels C and D plot the CPIE values for the real data as a function of turnover along with a dashed vertical line indicating the annual mean of daily turnover.

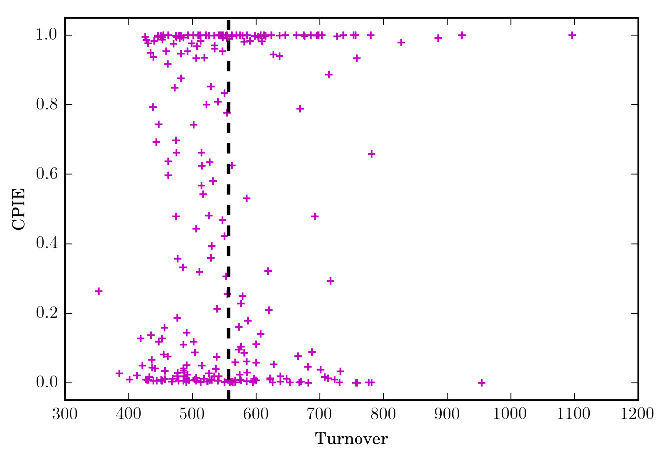
(a) XOM 1993



(b) XOM 2012



(c) XOM 1993



(d) XOM 2012

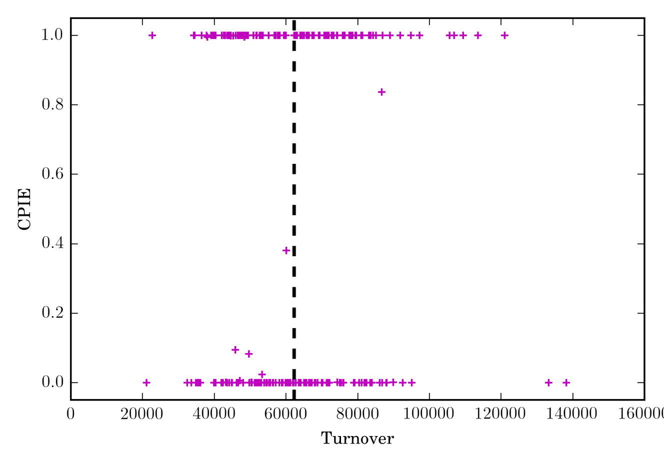


Figure 8: **OWR Tree.** In the OWR model, prior to markets opening, private information arrives with probability α. Once markets open, investors submit their trades generating order imbalance (y_e), and the intraday return (r_d). After markets close, private information becomes public and is reflected in the overnight return (r_o). The variables (y_e, r_d, r_o) are normally distributed with mean zero and covariance Σ, where Σ is function of the information arrival indicator (I). For instance, when there is no private information, there is a reversal in the returns ($cov(r_d, r_o) < 0$) and when there is private information there is a continuation in the returns ($cov(r_d, r_o) > 0$).

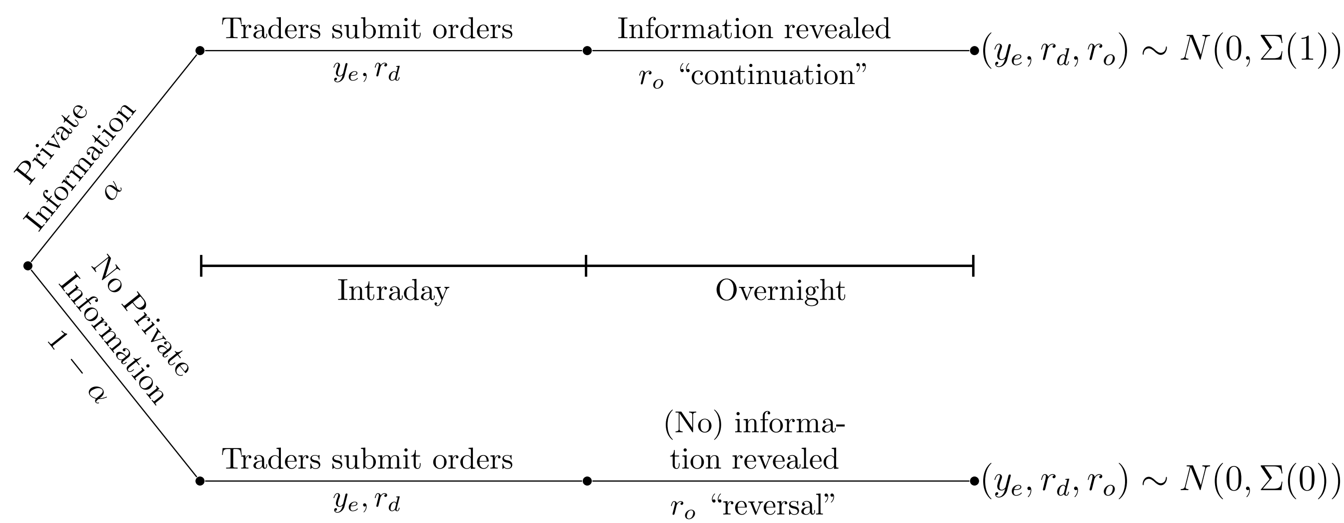


Figure 9: **OWR** α**.** This figure shows the distribution of yearly α parameter estimates for the OWR model. The solid black line represents the median value, and the dashed lines represent the 5, 25, 75, and 95 percentiles.

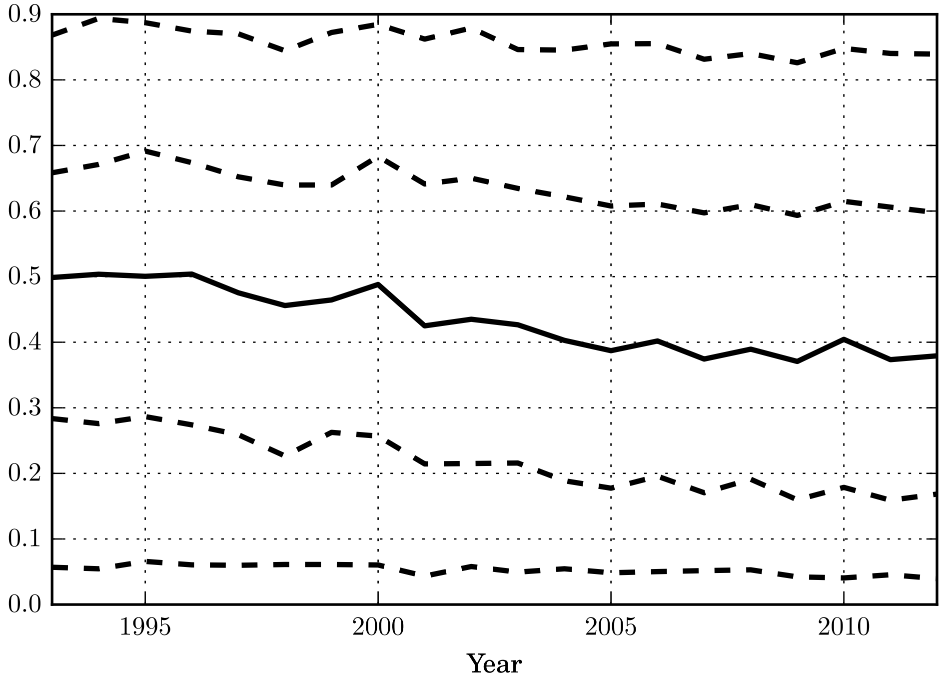
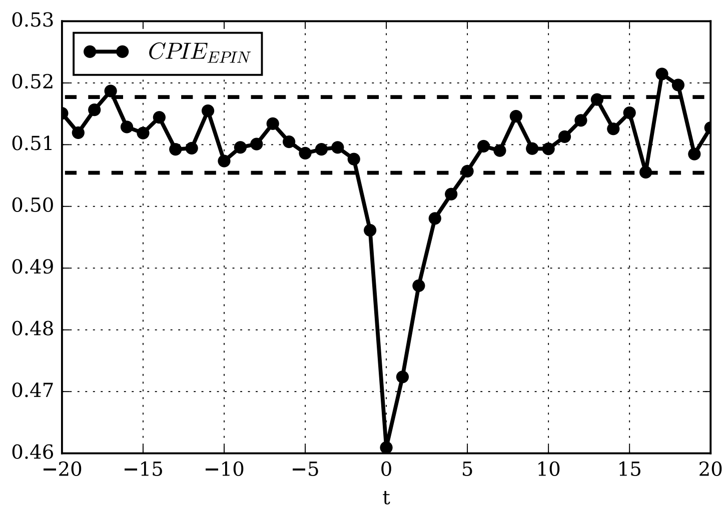


Figure 10: **Earnings Announcements.** Panel A (B) shows the average $CPIE_{EPIN}$ ($CPIE_{OWR}$) for the EPIN (OWR) model in event time surrounding earnings announcements.

(a) $CPIE_{EPIN}$



(b) $CPIE_{OWR}$

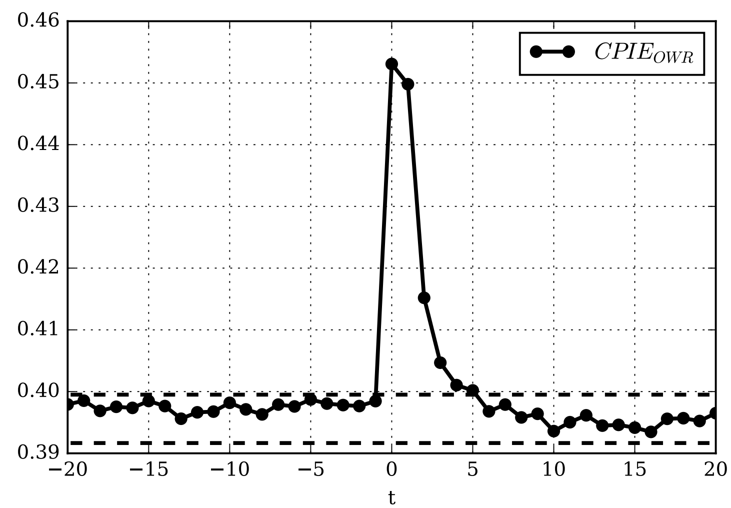
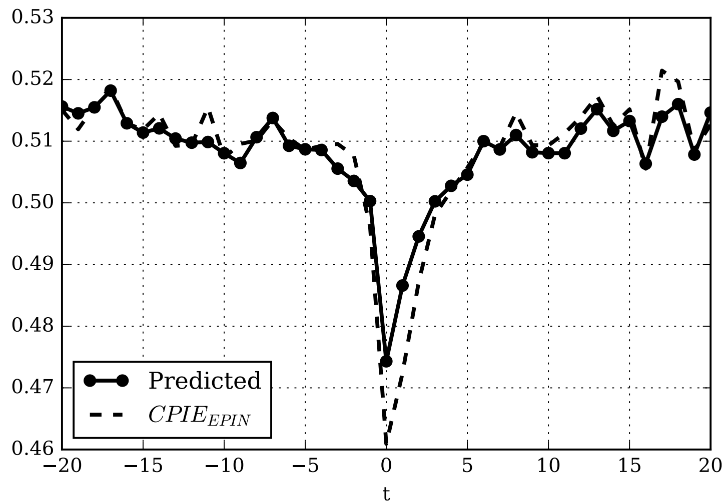


Figure 11: **Earnings Announcements - EPIN Decomposition.** Panels A and B compare the average $CPIE_{EPIN}$ with the $CPIE_{EPIN}$ predicted using either $\frac{|B-S|}{B+S}$ or turnover ($turn$), respectively. To obtain the predictions, we run regressions of daily $CPIE_{EPIN}$ on $\frac{|B-S|}{B+S}$ or $turn$, and their respective squared terms.

(a) Prediction using $\frac{|B-S|}{B+S}$ and $\frac{|B-S|}{B+S}^2$



(b) Prediction using $turn$ and $turn^2$

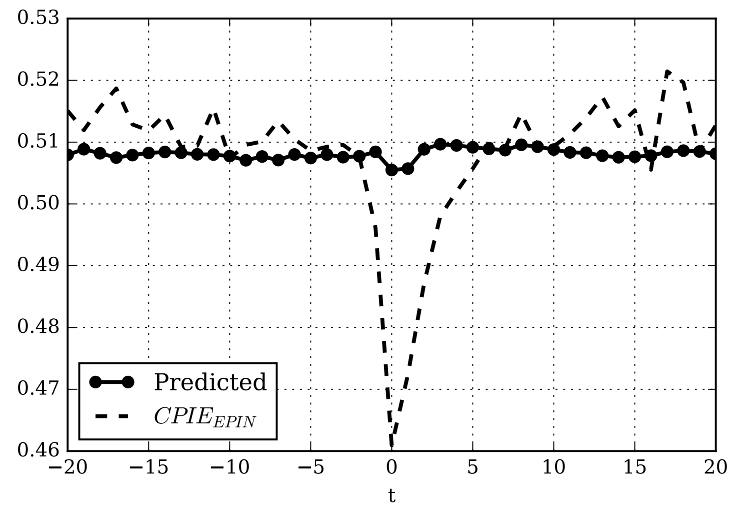
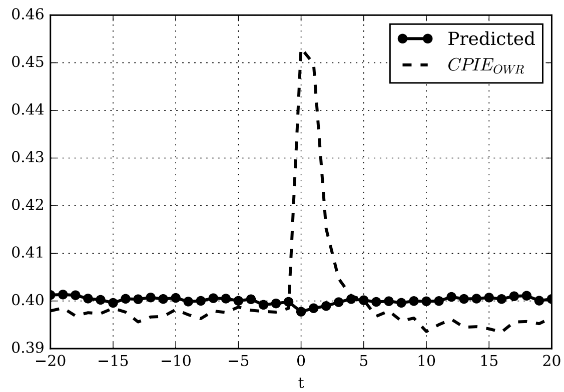
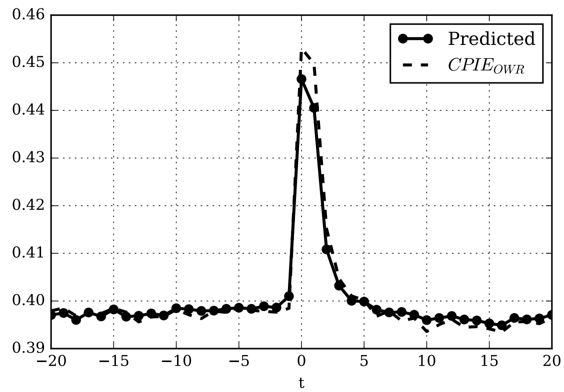


Figure 12: **Earnings Announcements - OWR Decomposition.** Panels A–F compare the average $CPIE_{OWR}$ with the $CPIE_{OWR}$ predicted using the squared and interaction terms of y_e, r_d, and r_o.

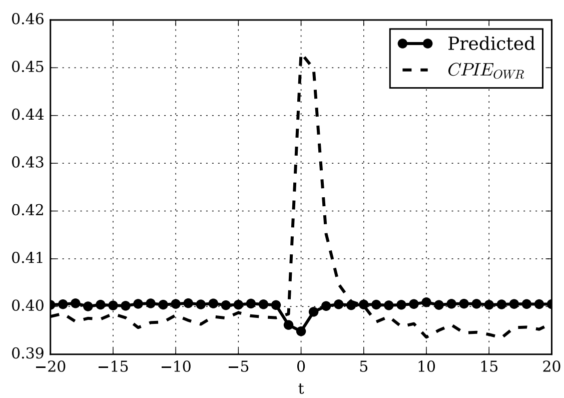
(a) Prediction using y_e^2



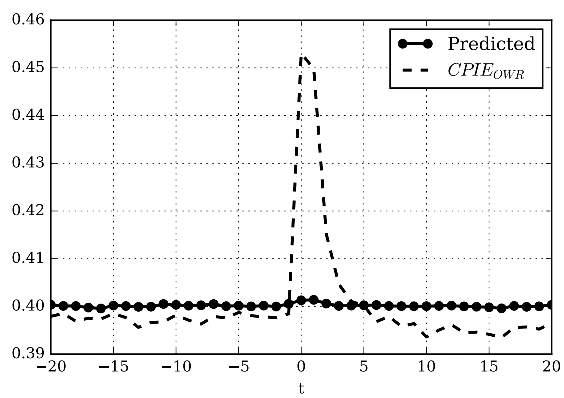
(b) Prediction using r_d^2



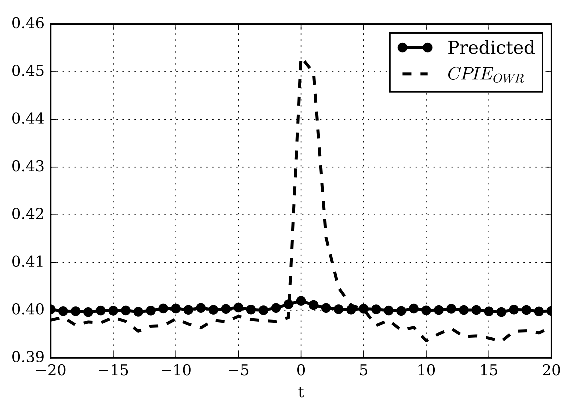
(c) Prediction using r_o^2



(d) Prediction using $y_e \times r_d$



(e) Prediction using $y_e \times r_o$



(f) Prediction using $r_d \times r_o$

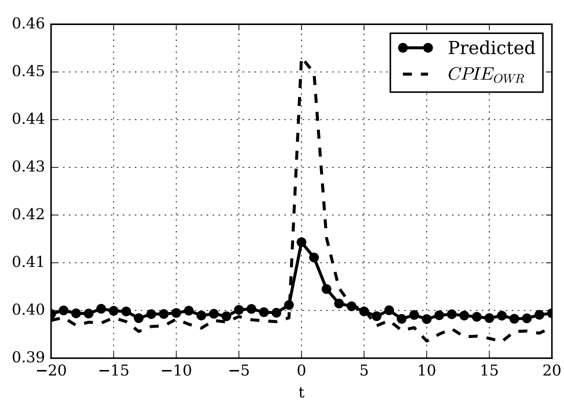
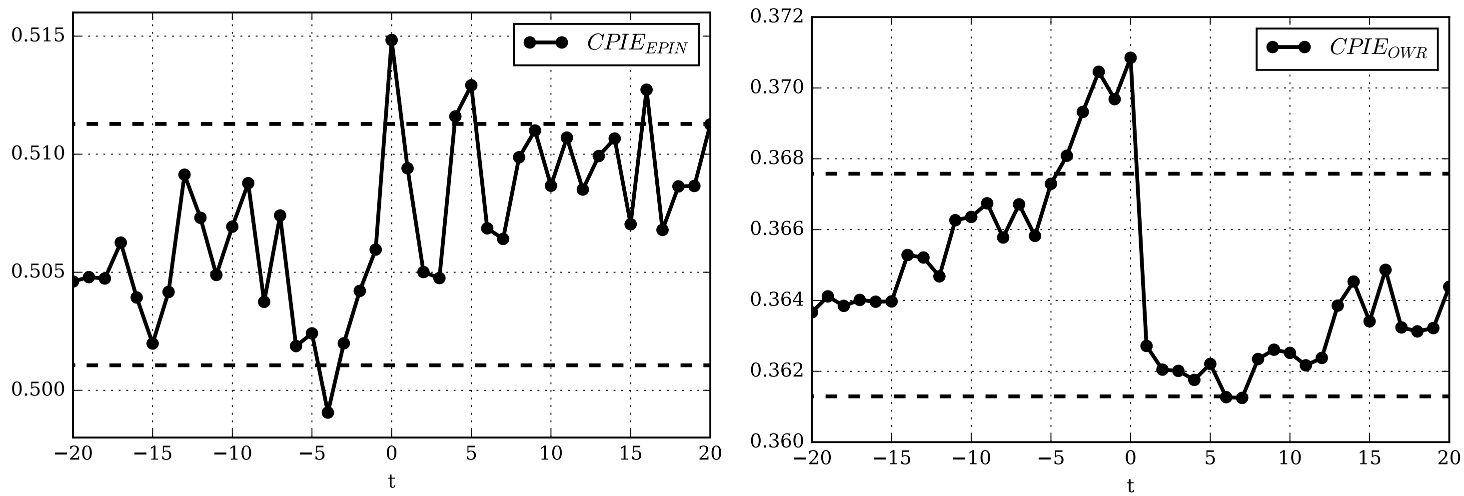


Figure 13: **Opportunistic Insider Trades.** Panel A (B) shows the average $CPIE_{EPIN}$ ($CPIE_{OWR}$) for the EPIN (OWR) model in event time surrounding opportunistic insider trades.

(a) $CPIE_{EPIN}$ (b) $CPIE_{OWR}$



Internet Appendix: What does the PIN model identify as private information?

Jefferson Duarte, Edwin Hu, and Lance Young

April 29$^{\text{th}}$, 2016

A The DY model

Duarte and Young (2009) propose an extension of the PIN model that accounts for the positive correlation between buys and sells. We show in this Appendix that the Duarte and Young (2009) model also performs poorly late in our sample from 1993–2012.

A.1 The DY model

Duarte and Young (2009) extend the PIN model to address some of its shortcomings in matching the order flow data. Specifically, the authors note that the PIN model implies that the number of buys and sells are negatively correlated; however, in the data the correlation between the number of buys and sells is overwhelmingly positive. To correct this problem, the DY model partially disentangles turnover variation from private information arrival. As in the PIN model, the DY model posits that at the beginning of each day, informed investors receive a private signal with probability α. If the private signal is positive, buy orders from the informed traders arrive according to a Poisson distribution with intensity μ_B. If the private signal is negative, informed sell orders arrive according to a Poisson distribution with intensity μ_S. If the informed traders receive no private signal, they do not trade.

In contrast to the PIN model, the DY model allows for symmetric order flow shocks. These shocks increase both the number of buyer- and seller-initiated trades but are unrelated to private information events. Symmetric order flow shocks can happen for a variety of reasons, such as disagreement among traders about the interpretation of public news. Alternatively, liquidity shocks may occur that cause investors holding different collections of assets to simultaneously rebalance their portfolios, resulting in increases to both buys and sells. Regardless of the mechanism, symmetric order flow shocks arrive on any given day with probability θ. On days with symmetric order flow shocks, both the number of buyer- and seller-initiated trades increase by amounts drawn from independent Poisson distributions with intensity Δ_B or Δ_S, respectively. Buy and sell orders from uninformed traders arrive according to a Poisson distribution with intensities ϵ_B ($\epsilon_B + \Delta_B$) and ϵ_S ($\epsilon_S + \Delta_S$) on days without (with) symmetric order flow shocks. Fig. A1 shows the structure of the DY model.

Under the DY model, turnover can increase due to either symmetric order flow shocks or the arrival of private information. To see this, note that the expected number of buys plus sells on days with positive (negative) information and without symmetric order flow shocks is $\epsilon_B + \epsilon_S + \mu_B$ ($\epsilon_B + \epsilon_S + \mu_S$); the expected number of trades on days with symmetric order flow shocks and without private information shocks is $\epsilon_B + \epsilon_S + \Delta_B + \Delta_S$, and the expected number of trades is $\epsilon_B + \epsilon_S$ on days without either.

A.2 Estimation of the DY model

As with the PIN model, we estimate the DY model numerically via maximum likelihood. Let $\Theta_{DY,i} = (\alpha_i, \mu_{B_i}, \mu_{S_i}, \epsilon_{B_i}, \epsilon_{S_i}, \delta_i, \theta_i, \Delta_{B_i}, \Delta_{S_i})$ be the vector of parameters of the DY model for stock i. Let $B_{i,t}$ and $S_{i,t}$ be the number of buys and sells, respectively, for stock i on day t. Let $D_{DY,i,t} = [B_{i,t}, S_{i,t}, \Theta_{DY,i}]$. The likelihood function of the extended model is $\prod_{t=1}^{T} L(D_{DY,i,t})$:

$$\begin{aligned} L(D_{DY,i,t}) &= L_{NI,NS}(D_{DY,i,t}) + L_{NI,S}(D_{DY,i,t}) + L_{I^-,NS}(D_{DY,i,t}) \\ &\quad + L_{I^-,S}(D_{DY,i,t}) + L_{I^+,NS}(D_{DY,i,t}) + L_{I^+,S}(D_{DY,i,t}) \end{aligned} \tag{1}$$

where $L_{NI,NS}(D_{DY,i,t})$ is the likelihood of observing $B_{i,t}$ and $S_{i,t}$ on a day without private information or a symmetric order flow shock; $L_{NI,S}(D_{DY,i,t})$ is the likelihood of $B_{i,t}$ and $S_{i,t}$ on a day without private information but with a symmetric order flow shock; $L_{I^-,NS}$ ($L_{I^-,S}$) is the likelihood of $B_{i,t}$ and $S_{i,t}$ on a day with negative information and without (with) a symmetric order flow shock; and $L_{I^+,NS}$ ($L_{I^+,S}$) is the probability on a day with positive information and without (with) a symmetric order flow shock. Analogous to the original PIN model, each term in the likelihood function corresponds to a branch in the tree in Fig. A1 and each term is given by:

$$L_{NI,NS}(D_{DY,i,t}) = (1-\alpha_i)(1-\theta_i)e^{-\epsilon_{B_i}}\frac{\epsilon_{B_i}^{B_{i,t}}}{B_{i,t}!}e^{-\epsilon_{S_i}}\frac{\epsilon_{S_i}^{S_{i,t}}}{S_{i,t}!} \tag{2}$$

$$L_{NI,S}(D_{DY,i,t}) = (1-\alpha_i)\theta_i e^{-(\epsilon_{B_i}+\Delta_{B_i})}\frac{(\epsilon_{B_i}+\Delta_{B_i})^{B_{i,t}}}{B_{i,t}!}e^{-(\epsilon_{S_i}+\Delta_{S_i})}\frac{(\epsilon_{S_i}+\Delta_{S_i})^{S_{i,t}}}{S_{i,t}!} \tag{3}$$

$$L_{I^-,NS}(D_{DY,i,t}) = \alpha_i(1-\theta_i)(1-\delta_i)e^{-\epsilon_{B_i}}\frac{\epsilon_{B_i}^{B_{i,t}}}{B_{i,t}!}e^{-(\mu_{S_i}+\epsilon_{S_i})}\frac{(\mu_{S_i}+\epsilon_{S_i})^{S_{i,t}}}{S_{i,t}!} \tag{4}$$

$$L_{I^-,S}(D_{DY,i,t}) = \alpha_i\theta_i(1-\delta_i)e^{-(\epsilon_{B_i}+\Delta_{B_i})}\frac{(\epsilon_{B_i}+\Delta_{B_i})^{B_{i,t}}}{B_{i,t}!}e^{-(\mu_{S_i}+\epsilon_{S_i}+\Delta_{S_i})}\frac{(\mu_{S_i}+\epsilon_{S_i}+\Delta_{S_i})^{S_{i,t}}}{S_{i,t}!} \tag{5}$$

$$L_{I^+,NS}(D_{DY,i,t}) = \alpha_i(1-\theta_i)\delta_i e^{-(\mu_{B_i}+\epsilon_{B_i})}\frac{(\mu_{B_i}+\epsilon_{B_i})^{B_{i,t}}}{B_{i,t}!}e^{-\epsilon_S}\frac{\epsilon_{S_i}^{S_{i,t}}}{S_{i,t}!} \tag{6}$$

$$L_{I^+,S}(D_{DY,i,t}) = \alpha_i\theta_i\delta_i e^{-(\mu_{B_i}+\epsilon_{B_i}+\Delta_{B_i})}\frac{(\mu_{B_i}+\epsilon_{B_i}+\Delta_{B_i})^{B_{i,t}}}{B_{i,t}!}e^{-(\epsilon_{S_i}+\Delta_{S_i})}\frac{(\epsilon_{S_i}+\Delta_{S_i})^{S_{i,t}}}{S_{i,t}!} \tag{7}$$

In order to avoid local optima, we use the maximum of the likelihood maximization with ten different starting points as in Duarte and Young (2009). In addition, for one of the starting points we choose (ϵ_B, ϵ_S) values, and $(\epsilon_B+\Delta_B, \epsilon_S+\Delta_S)$ equal to the sample means of buys and sells computed by the k-means algorithm with k=2. The k-means algorithm looks for clusters in the buys and sells such that each observation belongs to the cluster with the nearest mean. Because we know a priori that buys and sells have a strong positive correlation (see Duarte and Young (2009)), we partition the sample into high and low order flow clusters, which correspond to the symmetric order flow shock/no symmetric order flow shock states in the DY model. The other nine starting points are randomized. This procedure ensures that at least one of the starting points is centered properly, as the numerical likelihood estimation using purely random starts often stops at points outside of the central clusters of data.

A.3 $CPIE_{DY}$

As with the PIN model, for each stock-day, we compute the probability of an information event conditional on both the model parameters and on the number of buys and sells observed that day. Specifically, let the indicator $I_{i,t}$ take the value of one if an information event occurs for stock i on day t and zero otherwise. We compute $CPIE_{DY,i,t} = P\left[I_{i,t}=1|D_{DY,i,t}\right]$ as:

$$CPIE_{DY,i,t} = \frac{L_{I^+,NS}(D_{DY,i,t}) + L_{I^+,S}(D_{DY,i,t}) + L_{I^-,S}(D_{DY,i,t}) + L_{I^-,NS}(D_{DY,i,t})}{L(D_{DY,i,t})} \tag{8}$$

Analogous to the PIN model, the *Adj. PIN* of a stock is $\frac{\alpha(\delta\mu_B+(1-\delta)\mu_S)}{\alpha(\delta\mu_B+(1-\delta)\mu_S)+\varepsilon_B+\varepsilon_S+\theta(\Delta_B+\Delta_S)}$. This is the unconditional probability that any given trade is initiated by an informed trader. $CPIE_{DY}$ and *Adj.* PIN are linked via the unconditional probability of an information event, α, which is also the unconditional expectation of $CPIE_{DY}$.

Table A1 contains summary statistics for the parameter estimates for the DY model as well as summary statistics of the cross-sectional sample means and standard deviations of $CPIE_{DY}$. We see that the mean $CPIE$ behaves exactly like α. Hence, changes in $CPIE_{DY}$ and changes in the estimated alphas are analogous.

A.4 How does the DY model identify private information?

To illustrate how the $CPIE_{DY}$ works, we present a stylized example of the DY model in Fig. A2. In Panel A we plot simulated and real order flow data for Exxon-Mobil during 1993, with buys on the horizontal axis and sells on the vertical axis. Real data are marked as +, and simulated data as transparent dots. The real data are shaded according to the $CPIE$, with lighter points (+ cyan) representing low and darker points (+ magenta) high $CPIE$s.

The DY model generates six data clusters, greatly improving upon the PIN model's coverage of the data in 1993. The two clusters on the dotted line are not related to private information, but the other four clusters are. An econometrician using the DY model, moving along the dotted line, would observe that high turnover days–considered information days under the PIN model–are no longer classified as such, because higher turnover may be driven by symmetric order flow shocks under the DY model. Instead, the DY model identifies private information when moving away from the dotted line; when buys are greater than sells and vice versa.

Unfortunately, late in the sample the DY model breaks down. Panel B of Fig. A2 shows that the DY model, like the PIN model, fails to fit the majority of the order flow data for Exxon-Mobil in 2012. The problem of fitting the data is not limited to our stylized example. Fig. A3 shows that after 2005 the DY model estimates that the total likelihood for 80% of the order flow data of the median stock is less than 10^{-10}.

As a more formal test of the DY model, Table A2 presents regressions of $CPIE_{DY}$ based on simulated and real data. The right-hand side variables are the absolute order imbalance adjusted for buy/sell correlations ($|adj.OIB|$), turnover and its squared term. We define the adjusted absolute order imbalance as the absolute value of the residual from a regression of buys on sells. We use this measure to analyze the DY model because, as Fig. A2 suggests, the DY model implies that days with information events are far from the dashed line in this figure.[1] Turnover, as before, is defined as the sum of buys and sells. We report median coefficient estimates and $t-$statistics across all firms within a particular year. The coefficients are standardized as above. We report the average of the median, the 5^{th}, and the 95^{th} percentiles of the R^2s and R^2_{inc}s.

As with the $CPIE_{PIN}$, in theory, turnover has little additional power in explaining $CPIE_{DY}$. The incremental R^2s in Table A2 Panel A are low with an average value close to 4%. This is smaller than the average incremental R^2s of the PIN model. The intuition for this result is that the DY model disentangles turnover and order flow shocks by including the possibility of symmetric order flow shocks. Buying and selling activity can simultaneously be higher than average, but this is not indicative of private information unless there is a large order flow imbalance.

Panel B of Table A2 reports regression results for the real, rather than simulated, data. The DY model behaves very differently when using real data as opposed to data generated from the model. The R^2s for the real data are much lower than those in the simulated data, declining from 35% in 1993 to 12% in 2012. The incremental R^2 indicates that turnover and turnover squared explain a large degree of variation in $CPIE_{DY}$. Indeed, the average ratio of the median R^2s, $R^2_{inc.}/(R^2 + R^2_{inc.})$, is about 40%. The p-values are the average probability (under the DY model) of observing an incremental R^2 larger or equal to the observed in the real data and $\%Rej.$ is the frequency that we reject the null hypothesis that the incremental R^2 is consistent with the DY model at 5% significance. In 1993, our hypothesis test based rejects the model at 5% significance for 48% of the stocks, while in 2012 this percentage increases to around 70%.

[1] Our results are qualitatively similar if we use absolute order imbalance instead of adjusted absolute order imbalance.

B Estimating Order Flow, $r_{o,i,t}$ and $r_{d,i,t}$

Wharton Research Data Services (WRDS) provides trades matched to National Best Bid and Offer (NBBO) quotes at 0, 1, 2, and 5 second delay intervals. We use only "regular way" trades, with original time and/or corrected timestamps to avoid incorrect quotes or non-standard settlement terms. For instance, trades that are settled in cash or settled the next business day.[2] Prior to 2000, we match "regular way" trades to quotes delayed for 5 seconds; between 2000 and 2007, we match trades to quotes delayed for 1 second; and after 2007, we match trades to quotes without any delay.

We classify the matched trades as either buys or sells following the Lee and Ready (1991) algorithm, which classifies all trades occurring above (below) the bid-ask mid-point as buyer (seller) initiated. We use a tick test to classify trades that occur at the mid-point of the bid and ask prices. The tick test classifies trades as buyer (seller) initiated if the price was above/(below) that of the previous trade.

To estimate $r_{o,i,t}$ and $r_{d,i,t}$, we run daily cross-sectional regressions of overnight and intraday returns on a constant, historical β (based on the previous 5 years of monthly CRSP returns), log market cap, log book-to-market (following Fama and French (1992), Fama and French (1993), and Davis, Fama, and French (2000)). We impose min/max values for book equity (before taking logs) of 0.017 and 3.13, respectively. If book equity is negative, we set it to 1 before taking logs, so that it is zero after taking logs. We use the residuals from these daily cross-sectional regressions, winsorized at the 1 and 99% levels as our idiosyncratic intraday ($r_{d,i,t}$) and overnight ($r_{o,i,t}$) returns.

C Details of the PIN model

C.1 PIN Likelihood

Let $B_{i,t}$ ($S_{i,t}$) represent the number of buys (sells) for stock i on day t and $\Theta_{PIN,i} = (\alpha_i, \mu_i, \epsilon_{B_i}, \epsilon_{S_i}, \delta_i)$ represent the vector of the PIN model parameters for stock i. Let $D_{PIN,i,t} = [\Theta_{PIN,i}, B_{i,t}, S_{i,t}]$. The likelihood of observing $B_{i,t}$ and $S_{i,t}$ on a day without an information event, on a day with positive information event, and on a day with a negative

[2]Trade COND of ("@","*", or " ") and CORR of (0,1)

information event are:

$$L_{NI}(D_{PIN,i,t}) = (1-\alpha_i)e^{-\epsilon_{B_i}}\frac{\epsilon_{B_i}^{B_{i,t}}}{B_{i,t}!}e^{-\epsilon_{S_i}}\frac{\epsilon_{S_i}^{S_{i,t}}}{S_{i,t}!} \tag{9}$$

$$L_{I^+}(D_{PIN,i,t}) = \alpha_i\delta_i e^{-(\mu_i+\epsilon_{B_i})}\frac{(\mu_i+\epsilon_{B_i})^{B_{i,t}}}{B_{i,t}!}e^{-\epsilon_{S_i}}\frac{\epsilon_{S_i}^{S_{i,t}}}{S_{i,t}!} \tag{10}$$

$$L_{I^-}(D_{PIN,i,t}) = \alpha_i(1-\delta_i)e^{-\epsilon_{B_i}}\frac{\epsilon_{B_i}^{B_{i,t}}}{B_{i,t}!}e^{-(\mu_i+\epsilon_{i,S})}\frac{(\mu_i+\epsilon_{i,S})^{S_{i,t}}}{S_{i,t}!} \tag{11}$$

where $L_{NI}(D_{PIN,i,t})$ is the likelihood of observing $B_{i,t}$ and $S_{i,t}$ on a day without private information trading; L_{I^-} (L_{I^+}) is the likelihood of $B_{i,t}$ and $S_{i,t}$ on a day with negative (positive) information.

C.2 Maximum likelihood procedure

To estimate the PIN likelihood function, we use the maximum of the likelihood maximization with ten different starting points as in Duarte and Young (2009). We note, however, that late in the sample, the likelihood functions of the PIN are very close to zero. After 2006, the PIN model suggests that 90% of the observed daily order flows for the median stock have a near-zero probability (i.e. smaller than 10^{-10}) of occurring. This makes the estimation susceptible to local optima. To get around this problem, we choose one of our ten starting points to be such that the PIN model clusters are close to the observed mean of the number of buys and sells. Specifically, we choose ϵ_B and ϵ_S values equal to the sample means of buys and sells, α equal to 1%, and delta equal to the mean absolute value of order imbalance. The other nine starting points are randomized. We do this in order to ensure that at least one of the starting points is centered properly, as the numerical likelihood estimation using purely random starts often stops at points outside of the central cluster of data.

C.3 Computing $CPIE_{PIN}$

In Section 2 of the paper, we define the $CPIE$ as the ratio of the "news" likelihood functions to the sum total of the likelihood functions. In practice, there are many cases in the PIN model for which the data a near-zero probability of occurring, meaning $L(D_{PIN,i,t}) = L_{NI}(D_{PIN,i,t}) + L_{I^+}(D_{PIN,i,t}) + L_{I^-}(D_{PIN,i,t})$ is smaller than 10^{-10}. As a result the $CPIE$ ratio frequently results in a divide by zero error.

In order to compute $CPIE$ for these days, we "center" the likelihoods around the state with the highest log-likelihood before computing the $CPIE$. For example, consider the PIN model with:

$$L_{\max} \equiv \max\{L_{NI}, L_{I^+}, L_{I^-}\}, \tag{12}$$

$$\ell_{\max} \equiv \log(L_{\max}) \tag{13}$$

where ℓ represents the log of the corresponding likelihood function. We compute the centered versions of each of the likelihood functions:

$$\ell'_{NI} = \ell_{NI} - \ell_{\max}, \tag{14}$$

$$\ell'_{I^+} = \ell_{I^+} - \ell_{\max}, \tag{15}$$

$$\ell'_{I^-} = \ell_{I^-} - \ell_{\max}. \tag{16}$$

We compute the $CPIE'$ as:

$$CPIE'_{PIN} = \frac{L'_{I^+} + L'_{I^-}}{L'_{NI} + L'_{I^+} + L'_{I^-}} \tag{17}$$

such that the most likely state has $L' = 1$. For a high turnover day, it may be the case that $L'_{I^+} = 1$, $L'_{I^-} = 0$ and $L'_{NI} = 0$; hence, the $CPIE$' will be 1. This computational procedure is equivalent to taking the limit of $CPIE_{PIN}$ as $L(D_{PIN,i,t})$ goes to zero. We follow a similar procedure to compute $CPIE_{DY}$.

C.4 $CPIE_{PIN}$ of M&A targets around announcements

Aktas, de Bodt, Declerck, and Van Oppens (2007) find that PIN is higher after merger announcements than before, partially as a result of increases in PIN model's α. In this section we show that their results are related to our main finding that the PIN model identifies private information from turnover.

We examine the period $t \in [-30, 30]$ around the event. To do so, we estimate the parameter vector $\Theta_{PIN,i}$ in the period $t \in [-312, -60]$ before the event and then compute the daily $CPIE$s for the period $t \in [-30, 30]$ surrounding the announcement.

Panel A of Fig. A4 shows the average $CPIE_{PIN}$ in event time for our sample of M&A targets. The graph shows that, under the PIN model, the probability of an information event increases prior to the event, starting at around 55% 20 days before the announcement and peaking around 80% on the after day of the announcement. The rise in the probability of an information event prior to the announcement is consistent with a world where informed traders generate signals about potential mergers and acquisitions and trade on this information before the events are announced to the public. However, $CPIE_{PIN}$ is also higher *after* the actual announcements become public information. In fact, $CPIE_{PIN}$ remains above the average $CPIE_{PIN}$ observed in the gap period, $[-60, -31]$, for 20 trading days after the announcement.

Panels B and C of Fig. A4 shed light on the features of the data that produce the observed pattern in the average $CPIE_{PIN}$ in Panel A. Panel B shows the average predictions from OLS regressions of $CPIE_{PIN}$ on order imbalance and absolute order imbalance squared across all of the stocks in the event study sample. The solid line indicates that order imbalance explains only a small fraction of the movement in $CPIE_{PIN}$ during the event window. Panel C shows the average predictions from regressions of $CPIE_{PIN}$ on turnover and turnover squared. The solid line indicates that the variation in $CPIE_{PIN}$ around M&A announcements is explained almost entirely by turnover. The intuition follows directly from the main results, which illustrates that $CPIE_{PIN}$ is mechanically driven by turnover increases. The higher post-event turnover levels are enough to keep $CPIE_{PIN}$ above its pre-event mean for a substantial period.

D Details of the EPIN model

The EPIN model extends the PIN model to allow for continuous variation in turnover unrelated to private information arrival.

D.1 The microstructure of the EPIN model

The market maker knows that the number of trades (i.e. $B + S$) on day t is distributed as a Poisson random variable with intensity λ_t. The trade intensity, λ_t, is drawn from a Gamma distribution with parameters r and p. In what follows, in the interest of clarity, we suppress the t subscript on λ. The market maker does not observe λ directly, she only sees the buy and sell orders as they arrive. The market maker also knows that at the beginning of every day the probability that informed traders receive a private signal is α. If the informed receive a private signal, then the market maker knows that some fraction of the day's total number of trades will be informed. If the informed traders receive no private signal, then all trades are uninformed. If there is no information in the market, then conditional on λ, the sum of buys and sells is drawn from a Poisson distribution with arrival rate λ. If informed traders do receive a private signal, η represents the ratio of the expected number of informed to uninformed trades. Thus, if informed traders receive a private signal then the fraction of informed trade to total trade is $\frac{\eta}{1+\eta}$. The corresponding fraction of uninformed trade is equal to $1 - \frac{\eta}{1+\eta} = \frac{1}{1+\eta}$. Thus, if informed traders receive a private signal, then conditional on λ, the total arrival rate of orders remains equal to $(\frac{1}{1+\eta} + \frac{\eta}{1+\eta})\lambda{=}\lambda$. It is immediately clear from this intuition that the probability of informed trade under the EPIN model is simply the unconditional expected fraction of informed trade to total trade, $PIN_{EPIN} = \frac{\alpha\eta}{1+\eta}$. The PIN_{EPIN} does not involve λ because λ determines the overall intensity of trade, but not the split between informed and uninformed trade.

Formally, the probability that any given trade is informed is equal to the expected number of informed trades divided by the expected number of trades. This ratio is:

$$\frac{\text{E[Inf. Trades]}}{\text{E[Trades]}} = \frac{\text{E[E[Inf. Trades}|\lambda]]}{\text{E[E[Trades}|\lambda]]}. \tag{18}$$

The numerator for the $EPIN$ is $\text{E}[\alpha\delta\big(\frac{\eta}{1+\eta}\big)\lambda + \alpha(1-\delta)\big(\frac{\eta}{1+\eta}\big)\lambda]$, and the denominator is simply $\text{E}[\lambda]$. Simplifying we get that $PIN_{EPIN} = \frac{\alpha\eta}{1+\eta}$.

To see the connection between the PIN_{PIN} and PIN_{EPIN}, first note that we can write the formula for PIN_{PIN} using Equation 18. Using the reparameterization of the PIN model presented in Section 3.1, the numerator is $\alpha \times E[\text{Inf. Trades}|\lambda = \lambda(1)] + (1-\alpha) \times E[\text{Inf. Trades}|\lambda = \lambda(0)]$. The expected number of informed trades on days with private information ($\lambda = \lambda(1)$) in the PIN model is μ and zero otherwise, hence the numerator of Equation 18 reduces to $\alpha \times \mu$. Under the PIN model, the denominator of Equation 18 is $\alpha \times E[\text{Trades}|\lambda = \lambda(1)] + (1-\alpha) \times E[\text{Trades}|\lambda = \lambda(0)]$. The expected number of trades on days with private information ($\lambda = \lambda(1)$) in the PIN model is $\epsilon_B + \epsilon_S + \mu$ and $\epsilon_B + \epsilon_S$ otherwise. Hence the denominator of Equation 18 reduces to $\epsilon_B + \epsilon_S + \alpha \times \mu$, which leads to the formula $PIN_{PIN} = \frac{\alpha\mu}{\alpha\mu+\varepsilon_B+\varepsilon_S}$. Note that unlike the PIN_{PIN}, α does not appear in the denominator of the PIN_{EPIN}. This difference occurs because, in the PIN model, everything else equal, stocks with higher α have higher expected turnover. This relation has a direct impact on the denominator of Equation 18 and comes about because of the conflation of expected turnover and the arrival of private information in the PIN model (see Equation 1 in the paper). In the EPIN model, on the other hand, expected turnover (λ) is drawn independently of private information arrival. Hence, α has no effect on expected turnover and thus no place in the denominator of Equation 18.

Finally, to verify that the EPIN model captures the same microstructure intuition as the PIN model, consider the bid-ask spread under the EPIN model and the PIN model. Following similar logic to that in Easley, Keifer, O'Hara and Paperman (1996), the expression for the opening bid-ask spread under the EPIN model is the same as that under the PIN model:

$$\frac{\alpha\eta}{1+\eta} \times (\overline{V} - \underline{V}) = PIN_{EPIN} \times (\overline{V} - \underline{V}) \tag{19}$$

where $\overline{V}$ is the value of the firm conditional on good news and $\underline{V}$ represents the value of the firm conditional on bad news.

D.2 Negative binomial distribution in EPIN model

In the EPIN model, conditional on λ_t the distribution of turnover $(B + S)$ is $Poisson$ with intensity λ_t. Moreover, λ_t is drawn from $Gamma(r, p/(1-p))$ distribution. Hence, the probability that $B + S$ is equal to x in a given day is:

$$f(x; r, p) = \int_0^\infty \frac{\lambda^x}{x!}\lambda^{r-1}\frac{e^{-\lambda(1-p)/p}}{(\frac{p}{1-p})^r\Gamma(r)}d\lambda = \frac{(1-p)^r p^{-r}}{\Gamma(r)}p^{r+x}\Gamma(r+x) \tag{20}$$

which is the well known $Negative\ Binomial(r, p)$ (see Casella and Berger (2002)).

D.3 EPIN maximum likelihood estimation

Let $\Theta_{EPIN} = (\alpha, \delta, \eta, \theta, r, p)$ be the vector of parameters of the EPIN model. Let $B_{i,t}$ ($S_{i,t}$) represent the number of buys (sells) for stock i on day t and $D_{EPIN,i,t} = [\Theta_{EPIN,i}, B_{i,t}, S_{i,t}]$. The likelihood function of the extended PIN model is $\prod_{t=1}^T L(D_{EPIN,i,t})$, where

$$L(D_{PIN,i,t}) = L_{NI}(D_{EPIN,i,t}) + L_{I^+}(D_{EPIN,i,t}) + L_{I^-}(D_{EPIN,i,t}). \tag{21}$$

Define the function:

$$f(B,S;r,p,\theta) = \frac{\theta^B(1-\theta)^S}{B!S!}\frac{(1-p)^r p^{-r}}{\Gamma(r)} p^{r+B+S}\Gamma(r+B+S) \tag{22}$$

And the parameters $\theta_{I^+} = (\eta+\theta)/(1+\eta)$, $\theta_{I^-} = \theta/(1+\eta)$

$$\begin{aligned}
L_{NI}(D_{EPIN,i,t}) = && (1-\alpha)f(B,S;r,p,\theta) \\
L_{I^+}(D_{EPIN,i,t}) = && \alpha\delta f(B,S;r,p,\theta_{I^+}) \\
L_{I^+}(D_{EPIN,i,t}) = && \alpha(1-\delta)f(B,S;r,p,\theta_{I^-})
\end{aligned} \tag{23}$$

Conditional on λ_t and analogous to the original PIN model, each term in the likelihood function corresponds to a branch in the EPIN tree in the paper. We maximize the EPIN likelihood function in two steps. First we estimate the parameters r and p to fit the $Negative\ Binomial(r,p)$ distribution to the turnover data. We then maximize the EPIN likelihood with fixed r and p to obtain estimates of α, δ, η and θ. Analogous to the estimation of the PIN likelihood, in each step we use the maximum likelihood based on ten random starting points to avoid picking up local maxima.

D.4 Computing $CPIE_{EPIN}$

As with the PIN model, for each stock-day, we compute the probability of an information event conditional on both the model parameters and on the number of buys and sells observed that day. We compute $CPIE_{EPIN,i,t} = P\left[I_{i,t} = 1 | D_{DY,i,t}\right]$, which is equal to $(L_{I^-}(D_{EPIN,i,t}) + L_{I^+}(D_{EPIN,i,t}))/L(D_{EPIN,i,t})$. $CPIE_{EPIN}$ is:

$$CPIE_{EPIN} = \frac{\alpha\delta\theta_{I^+}^B(1-\theta_{I^+})^S + \alpha(1-\delta)\theta_{I^-}^B(1-\theta_{I^-})^S}{(1-\alpha)\theta^B(1-\theta)^S + \alpha\delta\theta_{I^+}^B(1-\theta_{I^+})^S + \alpha(1-\delta)\theta_{I^-}^B(1-\theta_{I^-})^S} \tag{24}$$

D.5 The EPIN model does not conflate turnover with private information

As a formal test of the EPIN model we run regressions of $CPIE_{EPIN}$ on the proportion of imbalanced trades $(\frac{|B-S|}{B+S})$ and a squared term $((\frac{|B-S|}{B+S})^2)$.[3] We use $\frac{|B-S|}{B+S}$ to analyze the

[3] We do not directly compare the simulations of the EPIN model to those of the PIN model. Instead we compare the real data for each model to the simulated data under the null hypothesis that each model

EPIN model because, as we discuss in the paper, the EPIN model implies that days with information events are the ones in which the proportion of imbalanced trades is large.

Panel A of Table A3 presents the results of regressions based on simulated data. As in the case of the regressions for the PIN model in the paper, we report the median coefficient estimates and t-statistics. The coefficients are standardized so they represent the increase in $CPIE_{EPIN}$ due to a one standard deviation increase in the corresponding independent variable. We also report the average of the median, the 5^{th}, and the 95^{th} percentiles of the empirical distribution of R^2s of these regressions generated by the 1,000 simulations. In general the EPIN model identifies private information from the proportion of imbalanced trades. The median R^2 values are high, ranging from 61%-92%, while the incremental R^2 from turnover is small-typically below 4%.

Panel B of Table A3 reports regression results for the real rather than simulated data. In contrast to the PIN model, in the real data the EPIN model identifies private information from the proportion of imbalanced trades and not turnover. The median R^2 values are high, ranging from 38%–72%, while the incremental R^2 from turnover is small—typically below 1%. Naturally, the EPIN model is not a perfect description of the order flow data. This can be seen from the fact that R^2 values using the real data are on average lower than those in the simulated data. However, the EPIN model fixes the conflation of arrival of private information with turnover, namely in the majority of stock-year observations in the real data the incremental R^2 due to turnover is at least as large as the incremental R^2 in the simulated data. Therefore, the EPIN model, while not a perfect description of the order flow data, fixes the problem of the PIN model which mechanically identifies private information from higher turnover.

E Details about the OWR model

E.1 OWR Likelihood

Let $\Theta_{OWR,i} = (\alpha_i, \sigma_{u_i}, \sigma_{z_i}, \sigma_{i_i}, \sigma_{p,d_i}, \sigma_{p,o_i})$ be the vector of parameters of this model. The parameter α_i is the probability that there is an information event on a given day. $\sigma^2_{z_i}$ is

identifies information consistent with the theory.

the variance of the noise of the observed net order flow (y_e); $\sigma^2_{u_i}$ is the variance of the net order flow from noise traders; $\sigma^2_{i_i}$ is the variance of the private signal received by the informed trader; σ^2_{p,d_i} is the variance of the intraday return; σ^2_{p,o_i} is the variance of the overnight return. Let $r_{d,i,t}$, ($r_{o,i,t}$) represent the intraday and overnight returns for stock i on day t, and ($y_{e,i,t}$) represent the order flow imbalance for stock i on day t. Let $D_{OWR,i,t} = [\Theta_{OWR,i}, r_{d,i,t}, r_{o,i,t}, y_{e,i,t}]$. The likelihood of observing $D_{OWR,i,t}$ on a day without and with an information event is:

$$L_{NI} = (1-\alpha) f_{NI}(D_{OWR,i,t}) \tag{25}$$

$$L_I = \alpha f_I(D_{OWR,i,t}) \tag{26}$$

where $f_{NI}(D_{OWR,i,t})$ is the joint probability density of $(y_{e,i,t}, r_{o,i,t}, r_{d,i,t})$ on days without information, $f_I(D_{OWR,i,t})$ is the density of $(y_{e,t}, r_{o,t}, r_{d,t})$ on days with information events. Both $f_{NI}(D_{OWR,i,t})$ and $f_I(D_{OWR,i,t})$ are multivariate normal with zero means and covariance matrices Ω_{NI_i} and Ω_{I_i}. The covariance matrix Ω_{NI_i} has elements:

$$Var(y_e) = \sigma^2_u + \sigma^2_z, \tag{27}$$

$$Var(r_d) = \sigma^2_{pd} + \alpha\sigma^2_i/4, \tag{28}$$

$$Var(r_o) = \sigma^2_{po} + \alpha\sigma^2_i/4, \tag{29}$$

$$Cov(r_d, r_o) = -\alpha\sigma^2_i/4, \tag{30}$$

$$Cov(r_d, y_e) = \alpha^{1/2}\sigma_i\sigma_u/2, \tag{31}$$

$$Cov(r_o, y_e) = -\alpha^{1/2}\sigma_i\sigma_u/2 \tag{32}$$

And Ω_{I_i}:

$$Var(y_e) = (1+1/\alpha)\sigma^2_u + \sigma^2_z, \tag{33}$$

$$Var(r_d) = \sigma^2_{pd} + (1+\alpha)\sigma^2_i/4, \tag{34}$$

$$Var(r_o) = \sigma^2_{po} + (1+\alpha)\sigma^2_i/4, \tag{35}$$

$$Cov(r_d, r_o) = (1-\alpha)\sigma^2_i/4, \tag{36}$$

$$Cov(r_d, y_e) = \alpha^{-1/2}\sigma_i\sigma_u/2 + \alpha^{1/2}\sigma_i\sigma_u/2, \tag{37}$$

$$Cov(r_o, y_e) = \alpha^{-1/2}\sigma_i\sigma_u/2 - \alpha^{1/2}\sigma_i\sigma_u/2 \tag{38}$$

E.2 How does the OWR model identify private information?

In theory, the OWR model identifies private information from the covariance matrix of the three variables in the model $(y_{e,i,t}, r_{o,i,t}, r_{d,i,t})$. To analyze the model, we run the regression of $CPIE_{OWR}$ on the squared and interaction terms of $(y_{e,i,t}, r_{o,i,t}, r_{d,i,t})$:

$$CPIE_{OWR,i,t} = \beta_0+\beta_1 y_{e,i,t}^2+\beta_2 r_{d,i,t}^2+\beta_3 r_{o,i,t}^2+\beta_4 y_{e,i,t} r_{d,i,t}+\beta_5 y_{e.i,t} r_{o,i,t}+\beta_6 r_{d,i,t} r_{o,i,t}+u_{i,t}. \quad (39)$$

Panel A of Table A4 presents median coefficient estimates, t-statistics, and three percentiles of R^2s across all firms within a particular year using simulated data. The results highlight the intuition behind the model. The probability of an information event on any given day is increasing in the square of intraday returns, the interaction between imbalance and intraday (or overnight) returns, and the interaction between intraday and overnight returns. The coefficient estimates on the square of the order imbalance and on the square of overnight returns are too small to be precisely measured. The high R^2s indicate that, practically speaking, the square of intraday returns, the interaction between intraday and overnight returns and the interaction between intraday returns and order flow imbalance are sufficient to explain a large part of the variation in $CPIE_{OWR}$.

Panel B of Table A4 shows the median coefficient estimates, t-statistics, and the results of the hypothesis tests based on R^2s across all firms within a particular year using real data. Unlike the PIN and DY models, the coefficient estimates are consistent across the simulated and real data. For instance in simulated data regressions in Panel A, 2008 is the only year in which y_e^2 is the most important term. In the real data regressions in Panel B, 2008 is also the only year in which y_e^2 is the most important term, indicating that the model matches the features of the data quite well, even for clear outliers like 2008. Furthermore, as with the simulated data regressions, the high median R^2s indicate that a large part of the variation in $CPIE_{OWR}$ is explained by the squared and interaction terms of $(y_{e,i,t}, r_{o,i,t}, r_{d,i,t})$ as implied by the model. The average across years of the R^2s in Panel B is about 83% and these R^2s increase over time, reaching 90% in 2012. Moreover, we reject the null hypothesis that the R^2s observed in the real data are consistent with the OWR model at 5% level for about 40% of the sample in 1993 and for about 8% of the sample in 2012.

The high R^2s in Panel B imply that, in principle, any variable unrelated to private information under the OWR model has only a small incremental value in explaining the $CPIE_{OWR}$. To see this note that the typical R^2 in Panel B is around 85%. This suggests that any additional regressor, even if it explained 100% of the residual variation in the regressions in Panel B, could only marginally improve the R^2 from 85% to 100%. Note that in the case of the PIN and DY models, our results show that turnover, which in principle is a poor measure of private information, largely drives the PIN and DY models' identification of private information. In contrast, under the OWR model the variables related to private information in the model (squares and interactions of y_e, r_o, and r_d) can explain a fairly large amount of the variation in $CPIE_{OWR}$. As a result, any variable that is not related to private information in the OWR model can only explain a relatively small fraction of the variation in $CPIE_{OWR}$.

Table A1: **DY Estimates.** This table summarizes parameter estimates of the DY model for 21,206 `PERMNO`-Year samples from 1993–2012. α represents the average unconditional probability of an information event at the daily level. ϵ_B and ϵ_S represent the expected number of daily buys and sells given no private information or symmetric order flow shocks. μ_b, and μ_s represent the expected additional order flows given an information event, which is good news with probability δ and bad news with probability $1-\delta$. A symmetric order flow shock occurs with probability θ, in which case the expected number of buys and sells increase by Δ_B and Δ_S, respectively. $\overline{CPIE}$ and Std($CPIE$) are the `PERMNO`-Year mean and standard deviation of $CPIE_{DY}$.

	N	Mean	Std	Q1	Median	Q3
α	21,206	0.456	0.092	0.409	0.464	0.509
δ	21,206	0.550	0.192	0.441	0.541	0.680
θ	21,206	0.249	0.137	0.149	0.253	0.344
ϵ_b	21,206	1,418	4,571	26	158	866
ϵ_s	21,206	1,397	4,570	28	148	807
Δ_b	21,206	2,148	10,058	41	190	989
Δ_s	21,206	2,097	9,934	34	160	908
μ_b	21,206	290	575	29	119	310
μ_s	21,206	284	574	27	107	302
$\overline{CPIE}$	21,206	0.455	0.092	0.409	0.461	0.506
Std($CPIE$)	21,206	0.454	0.056	0.431	0.479	0.493

Table A2: **DY Model Regressions.** This table reports real and simulated regressions of the $CPIE_{DY}$ on absolute adjusted order imbalance (|adj. OIB|), and absolute adjusted order imbalance squared (|adj. OIB|^2). In Panel A, we simulate 1,000 instances of the PIN model for each `PERMNO`-Year in our sample (1993–2012) and report mean standardized estimates for the median stock, along with 5%, 50%, and 95% values of the R^2 ($R^2_{inc.}$) values. We compute the incremental $R^2_{inc.}$ as the R^2 attributed to $turn$ and $turn^2$ in an extended regression model. In Panel B, we report standardized estimates for the median stock using real data, along with the median R^2 and $R^2_{inc.}$ values, and tests of the null hypothesis that the observed relation between $CPIE_{DY}$ and $turn$ is consistent with the DY model. The p-value is the average probability of observing an $R^2_{inc.}$ at least as large as what is observed in the real data. The % Rej. is the fraction of stocks for which we reject the hypothesis at the 5% level.

(a) Simulated Data

	β		t		R^2			$R^2_{inc.}$		
	|adj. OIB|	|adj. OIB|^2	|adj. OIB|	|adj. OIB|^2	5%	50%	95%	5%	50%	95%
1993	0.518	-0.230	(10.88)	(-4.74)	52.28%	59.44%	66.01%	5.55%	9.86%	15.29%
1994	0.484	-0.214	(10.47)	(-4.42)	50.66%	58.06%	64.97%	5.56%	9.46%	14.95%
1995	0.475	-0.214	(9.96)	(-4.32)	46.81%	54.46%	61.69%	7.01%	11.71%	17.54%
1996	0.516	-0.229	(10.54)	(-4.60)	51.36%	58.62%	65.21%	5.18%	9.09%	14.31%
1997	0.513	-0.221	(10.33)	(-4.40)	50.55%	57.80%	64.50%	4.78%	8.57%	14.03%
1998	0.537	-0.236	(10.60)	(-4.49)	52.85%	60.14%	66.63%	4.00%	7.45%	12.31%
1999	0.607	-0.281	(11.92)	(-5.45)	56.53%	63.49%	69.68%	3.07%	6.11%	10.47%
2000	0.597	-0.272	(11.43)	(-5.09)	55.69%	62.59%	69.09%	2.82%	5.65%	9.73%
2001	0.729	-0.350	(13.81)	(-6.75)	65.81%	71.48%	76.83%	0.62%	1.87%	4.09%
2002	0.769	-0.371	(15.03)	(-7.28)	71.90%	76.37%	80.55%	0.24%	1.04%	2.41%
2003	0.805	-0.394	(16.06)	(-7.99)	74.77%	78.95%	82.78%	0.34%	1.19%	2.71%
2004	0.798	-0.385	(15.94)	(-7.61)	77.39%	81.40%	84.70%	0.23%	0.95%	2.22%
2005	0.787	-0.365	(16.23)	(-7.40)	79.40%	83.08%	86.23%	0.25%	0.97%	2.20%
2006	0.761	-0.332	(15.52)	(-6.74)	79.38%	83.00%	86.15%	0.45%	1.41%	2.88%
2007	0.736	-0.311	(12.97)	(-5.97)	69.81%	74.50%	79.19%	1.23%	2.93%	5.99%
2008	0.755	-0.317	(15.14)	(-6.52)	77.82%	81.67%	85.36%	0.34%	1.21%	2.82%
2009	0.768	-0.331	(16.09)	(-7.01)	79.54%	83.16%	86.38%	0.63%	1.70%	3.51%
2010	0.769	-0.329	(15.95)	(-7.01)	78.65%	82.63%	86.22%	0.56%	1.64%	3.66%
2011	0.754	-0.313	(15.47)	(-6.73)	77.75%	81.79%	85.71%	0.63%	1.87%	4.10%
2012	0.763	-0.328	(15.65)	(-7.01)	77.64%	81.93%	85.61%	0.89%	2.25%	4.69%

Table A2: **DY Model Regressions.** Continued.

(b) Real Data

	β		t		R^2	$R^2_{inc.}$		
	|adj. OIB|	|adj. OIB|^2	|adj. OIB|	|adj. OIB|^2	50%	50%	p-value	% Rej.
1993	0.369	-0.170	(7.61)	(-3.48)	34.07%	15.22%	23.83%	48.21%
1994	0.348	-0.150	(7.51)	(-3.16)	33.55%	14.53%	23.87%	48.38%
1995	0.342	-0.149	(6.99)	(-3.00)	30.15%	15.63%	29.41%	43.47%
1996	0.358	-0.164	(7.33)	(-3.42)	31.11%	14.19%	25.56%	50.64%
1997	0.334	-0.140	(6.49)	(-2.78)	28.00%	13.92%	26.26%	50.56%
1998	0.329	-0.136	(6.21)	(-2.62)	26.26%	12.97%	22.18%	57.16%
1999	0.365	-0.166	(6.91)	(-3.16)	27.89%	12.56%	18.93%	62.38%
2000	0.333	-0.145	(5.75)	(-2.55)	23.49%	11.88%	20.82%	62.06%
2001	0.374	-0.176	(6.38)	(-3.06)	25.25%	9.07%	15.71%	74.29%
2002	0.328	-0.130	(4.82)	(-1.90)	21.31%	9.08%	10.15%	82.14%
2003	0.334	-0.135	(4.84)	(-1.98)	21.55%	8.58%	10.51%	81.42%
2004	0.295	-0.104	(4.15)	(-1.46)	18.31%	9.57%	10.09%	83.63%
2005	0.279	-0.103	(4.03)	(-1.51)	16.23%	10.61%	11.10%	82.60%
2006	0.243	-0.083	(3.40)	(-1.17)	12.46%	11.15%	16.81%	77.86%
2007	0.219	-0.086	(3.14)	(-1.25)	9.66%	12.26%	25.72%	65.76%
2008	0.217	-0.086	(3.05)	(-1.23)	8.83%	11.92%	19.43%	74.90%
2009	0.230	-0.093	(3.24)	(-1.30)	10.04%	11.43%	19.40%	74.53%
2010	0.241	-0.103	(3.41)	(-1.49)	10.59%	12.38%	21.74%	71.55%
2011	0.245	-0.102	(3.45)	(-1.50)	10.35%	13.05%	21.61%	71.57%
2012	0.275	-0.127	(4.04)	(-1.86)	12.22%	12.20%	23.56%	70.88%

Table A3: **EPIN Model Regressions.** This table reports real and simulated regressions of the $CPIE_{EPIN}$ on the proportion of imbalanced trades $\left(\frac{|B-S|}{B+S}\right)$ and its square. In Panel A, we simulate 1,000 instances of the EPIN model for each `PERMNO`-Year in our sample (1993–2012) and report mean standardized estimates for the median stock, along with 5%, 50%, and 95% values of the R^2 ($R^2_{inc.}$) values. We compute the incremental $R^2_{inc.}$ as the R^2 attributed to *turn* and $turn^2$ in an extended regression model. In Panel B, we report standardized estimates for the median stock using real data, along with the median R^2 and $R^2_{inc.}$ values, and tests of the null hypothesis that the observed relation between $CPIE_{EPIN}$ and *turn* is consistent with the EPIN model. The p-value is the average probability of observing an $R^2_{inc.}$ at least as large as what is observed in the real data. The % Rej. is the fraction of stocks for which we reject the hypothesis at the 5% level.

(a) Simulated Data

	β		t		R^2			$R^2_{inc.}$		
	$\frac{\lvert B-S\rvert}{B+S}$	$\left(\frac{\lvert B-S\rvert}{B+S}\right)^2$	$\frac{\lvert B-S\rvert}{B+S}$	$\left(\frac{\lvert B-S\rvert}{B+S}\right)^2$	5%	50%	95%	5%	50%	95%
1993	0.382	-0.134	(8.22)	(-3.04)	57.61%	63.37%	68.65%	1.79%	4.07%	7.31%
1994	0.355	-0.119	(8.19)	(-2.83)	56.90%	62.64%	67.90%	1.76%	4.23%	7.74%
1995	0.350	-0.113	(7.86)	(-2.59)	59.18%	64.87%	69.82%	1.68%	3.86%	7.24%
1996	0.364	-0.122	(8.31)	(-2.90)	60.59%	65.85%	70.81%	1.60%	3.84%	6.94%
1997	0.369	-0.126	(8.03)	(-2.84)	58.63%	64.01%	69.13%	1.29%	3.34%	6.34%
1998	0.388	-0.131	(8.93)	(-3.03)	60.99%	66.95%	71.74%	1.02%	2.81%	5.69%
1999	0.465	-0.190	(10.90)	(-4.26)	64.29%	69.23%	73.64%	1.01%	2.71%	5.10%
2000	0.447	-0.171	(9.34)	(-3.60)	60.81%	65.74%	70.43%	0.82%	2.42%	4.95%
2001	0.425	-0.123	(6.77)	(-2.08)	59.82%	65.02%	70.21%	0.71%	2.13%	4.40%
2002	0.243	0.007	(2.86)	(0.08)	55.43%	61.22%	66.58%	0.52%	1.87%	3.97%
2003	0.033	0.202	(0.30)	(1.95)	56.10%	62.06%	67.76%	0.51%	1.78%	4.05%
2004	-0.477	0.679	(-4.25)	(6.10)	56.37%	62.52%	68.15%	0.38%	1.47%	3.43%
2005	0.343	-0.062	(3.38)	(-0.67)	64.83%	70.03%	74.47%	0.16%	0.86%	2.23%
2006	0.294	-0.018	(3.16)	(-0.21)	72.38%	77.14%	80.90%	0.06%	0.42%	1.30%
2007	0.778	-0.338	(17.81)	(-7.59)	86.47%	88.49%	90.35%	0.02%	0.17%	0.54%
2008	0.784	-0.335	(18.60)	(-7.90)	90.29%	91.75%	93.13%	0.01%	0.12%	0.42%
2009	0.774	-0.321	(19.72)	(-8.04)	91.13%	92.47%	93.73%	0.01%	0.12%	0.40%
2010	0.773	-0.318	(19.47)	(-7.97)	90.93%	92.27%	93.57%	0.01%	0.13%	0.45%
2011	0.783	-0.335	(19.80)	(-8.16)	91.08%	92.48%	93.67%	0.01%	0.11%	0.40%
2012	0.781	-0.332	(19.89)	(-8.23)	90.82%	92.27%	93.54%	0.01%	0.12%	0.41%

Table A3: **EPIN Model Regressions.** Continued.

(b) Real Data

	β		t		R^2	$R^2_{inc.}$		
	$\frac{\lvert B-S\rvert}{B+S}$	$\left(\frac{\lvert B-S\rvert}{B+S}\right)^2$	$\frac{\lvert B-S\rvert}{B+S}$	$\left(\frac{\lvert B-S\rvert}{B+S}\right)^2$	50%	50%	p-value	% Rej.
1993	0.336	-0.113	(8.20)	(-2.93)	57.90%	1.00%	87.77%	3.26%
1994	0.321	-0.108	(8.12)	(-2.92)	56.55%	1.11%	84.63%	3.30%
1995	0.317	-0.098	(7.99)	(-2.62)	58.03%	1.08%	82.66%	4.03%
1996	0.339	-0.117	(8.73)	(-3.06)	59.28%	0.99%	84.95%	3.08%
1997	0.339	-0.117	(8.38)	(-2.98)	57.53%	1.03%	82.25%	4.16%
1998	0.362	-0.132	(9.59)	(-3.34)	61.34%	0.88%	82.57%	3.73%
1999	0.433	-0.183	(11.55)	(-4.88)	62.95%	0.80%	81.82%	5.18%
2000	0.419	-0.168	(9.74)	(-3.95)	58.88%	0.75%	81.03%	4.00%
2001	0.402	-0.143	(7.32)	(-2.62)	50.55%	0.48%	84.33%	3.52%
2002	0.255	-0.020	(3.57)	(-0.27)	42.07%	0.47%	80.50%	3.75%
2003	0.126	0.101	(1.70)	(1.36)	40.55%	0.46%	80.20%	3.19%
2004	-0.067	0.280	(-0.88)	(3.54)	38.32%	0.42%	75.32%	4.72%
2005	0.249	-0.015	(3.29)	(-0.20)	41.68%	0.41%	70.49%	6.64%
2006	0.264	-0.021	(3.81)	(-0.34)	43.41%	0.36%	59.43%	13.40%
2007	0.762	-0.447	(16.12)	(-9.57)	66.36%	0.31%	40.73%	25.49%
2008	0.800	-0.480	(18.60)	(-11.20)	70.98%	0.23%	39.63%	26.42%
2009	0.813	-0.492	(19.08)	(-11.49)	71.79%	0.23%	39.13%	31.68%
2010	0.814	-0.488	(18.94)	(-11.44)	72.77%	0.21%	41.33%	28.77%
2011	0.809	-0.480	(18.79)	(-11.21)	71.67%	0.22%	39.58%	29.71%
2012	0.804	-0.475	(18.83)	(-11.14)	72.72%	0.20%	42.12%	26.87%

Table A4: **OWR Model Regressions.** This table reports real and simulated regressions of the $CPIE_{OWR}$ on the squared and interaction terms of y_e, r_d, and r_o. In Panel A, we simulate 1,000 instances of the OWR model for each `PERMNO`-Year in our sample (1993–2012) and report mean standardized estimates for the median stock, along with 5%, 50%, and 95% values of the R^2 values. In Panel B, we report standardized estimates for the median stock using real data, along with the median R^2 values, and tests of the null that the model fits the data. The p-value is the average probability of observing an R^2 at least as small as what is observed in the real data. The % Rej. is the fraction of stocks for which we reject the null at the 5% level.

(a) Simulated Data

	β						t						R^2		
	y_e^2	$y_e \times r_d$	$y_e \times r_o$	r_d^2	$r_d \times r_o$	r_o^2	y_e^2	$y_e \times r_d$	$y_e \times r_o$	r_d^2	$r_d \times r_o$	r_o^2	5%	50%	95%
1993	0.002	0.068	-0.003	0.017	0.016	0.096	(0.42)	(11.52)	(-0.66)	(2.71)	(3.34)	(17.78)	68.29%	79.86%	88.22%
1994	0.002	0.065	-0.003	0.018	0.017	0.093	(0.53)	(12.10)	(-0.67)	(3.14)	(3.80)	(18.95)	70.03%	81.70%	89.67%
1995	0.003	0.065	-0.003	0.019	0.018	0.093	(0.57)	(12.03)	(-0.71)	(3.14)	(4.00)	(18.83)	69.82%	81.98%	89.91%
1996	0.003	0.066	-0.003	0.020	0.019	0.094	(0.68)	(12.73)	(-0.76)	(3.77)	(4.43)	(20.14)	72.12%	83.64%	91.18%
1997	0.003	0.063	-0.003	0.018	0.018	0.092	(0.77)	(14.31)	(-0.80)	(4.05)	(4.73)	(21.45)	73.01%	85.04%	92.43%
1998	0.002	0.070	-0.004	0.018	0.017	0.102	(0.67)	(16.25)	(-1.01)	(4.14)	(4.70)	(24.53)	74.91%	86.68%	93.93%
1999	0.003	0.060	-0.003	0.017	0.018	0.093	(0.74)	(13.90)	(-0.75)	(3.88)	(4.86)	(22.15)	72.82%	84.70%	92.22%
2000	0.003	0.051	-0.002	0.017	0.019	0.085	(0.87)	(13.37)	(-0.58)	(4.20)	(5.64)	(22.86)	73.87%	85.03%	92.21%
2001	0.002	0.066	-0.004	0.014	0.014	0.098	(0.51)	(17.18)	(-1.15)	(3.72)	(4.25)	(26.22)	76.05%	87.58%	94.14%
2002	0.001	0.066	-0.003	0.012	0.013	0.099	(0.44)	(18.37)	(-1.03)	(3.40)	(3.89)	(27.41)	76.47%	87.94%	94.40%
2003	0.002	0.071	-0.005	0.014	0.013	0.105	(0.48)	(19.18)	(-1.53)	(3.50)	(3.84)	(27.86)	77.31%	88.81%	94.93%
2004	0.001	0.068	-0.005	0.012	0.012	0.100	(0.49)	(21.61)	(-1.91)	(4.05)	(4.06)	(30.04)	79.32%	90.05%	95.22%
2005	0.002	0.061	-0.005	0.012	0.012	0.086	(0.60)	(22.68)	(-2.02)	(4.35)	(4.35)	(31.06)	80.89%	90.80%	95.18%
2006	0.001	0.063	-0.004	0.011	0.011	0.089	(0.52)	(22.88)	(-1.91)	(3.95)	(4.14)	(30.37)	80.34%	90.48%	95.19%
2007	0.001	0.051	-0.003	0.002	0.004	0.068	(0.65)	(22.32)	(-1.69)	(0.78)	(1.68)	(28.67)	81.21%	90.63%	95.41%
2008	0.076	0.000	-0.001	0.000	0.004	0.001	(27.51)	(0.07)	(-0.25)	(0.10)	(1.42)	(0.29)	76.59%	88.91%	95.17%
2009	0.002	0.039	-0.002	0.001	0.005	0.060	(1.18)	(18.30)	(-0.73)	(0.35)	(2.36)	(27.24)	80.66%	90.07%	95.06%
2010	0.002	0.038	-0.002	0.000	0.000	0.046	(0.94)	(18.05)	(-1.34)	(0.13)	(0.23)	(22.24)	78.97%	88.62%	94.54%
2011	0.001	0.042	-0.002	0.000	0.000	0.055	(0.79)	(19.58)	(-1.37)	(0.11)	(0.16)	(24.64)	80.82%	90.39%	95.10%
2012	0.001	0.046	-0.003	0.000	0.000	0.055	(0.68)	(19.47)	(-1.55)	(0.11)	(0.22)	(23.02)	79.83%	89.47%	94.62%

Table A4: **OWR Model Regressions.** Continued.

(b) Real Data

	β						t						R^2
	y_e^2	$y_e \times r_d$	$y_e \times r_o$	r_d^2	$r_d \times r_o$	r_o^2	y_e^2	$y_e \times r_d$	$y_e \times r_o$	r_d^2	$r_d \times r_o$	r_o^2	50%
1993	-0.000	0.053	-0.000	0.032	0.029	0.055	(-0.03)	(7.24)	(-0.13)	(4.41)	(4.56)	(8.11)	69.97%
1994	0.000	0.053	-0.001	0.032	0.027	0.060	(0.06)	(8.11)	(-0.17)	(4.69)	(4.68)	(9.44)	72.00%
1995	0.001	0.052	-0.001	0.033	0.029	0.059	(0.15)	(7.92)	(-0.17)	(4.74)	(4.89)	(9.35)	72.73%
1996	0.001	0.055	-0.003	0.032	0.028	0.062	(0.28)	(8.61)	(-0.52)	(4.77)	(4.81)	(9.83)	73.65%
1997	0.002	0.054	-0.002	0.029	0.027	0.061	(0.36)	(8.90)	(-0.53)	(4.85)	(4.84)	(10.17)	74.72%
1998	0.002	0.069	-0.004	0.025	0.023	0.074	(0.37)	(11.25)	(-0.89)	(4.43)	(4.15)	(12.61)	77.46%
1999	0.002	0.057	-0.003	0.025	0.025	0.065	(0.56)	(9.59)	(-0.64)	(4.33)	(4.58)	(11.66)	76.48%
2000	0.003	0.050	-0.003	0.021	0.022	0.066	(0.82)	(10.58)	(-0.98)	(4.50)	(5.15)	(14.37)	79.83%
2001	0.001	0.068	-0.003	0.018	0.016	0.078	(0.47)	(14.62)	(-0.94)	(4.10)	(3.81)	(16.91)	83.25%
2002	0.002	0.072	-0.002	0.016	0.014	0.081	(0.47)	(16.83)	(-0.72)	(3.88)	(3.71)	(19.17)	84.71%
2003	0.002	0.080	-0.003	0.017	0.015	0.080	(0.60)	(20.66)	(-0.94)	(4.38)	(3.93)	(20.51)	87.22%
2004	0.001	0.077	-0.005	0.016	0.012	0.074	(0.54)	(24.74)	(-1.74)	(4.48)	(3.58)	(21.11)	88.70%
2005	0.002	0.072	-0.005	0.013	0.010	0.065	(0.83)	(25.08)	(-2.12)	(4.36)	(3.32)	(20.58)	89.54%
2006	0.002	0.072	-0.005	0.013	0.010	0.066	(0.74)	(25.53)	(-1.61)	(4.12)	(3.36)	(20.42)	89.47%
2007	0.002	0.058	-0.003	0.004	0.005	0.058	(0.98)	(18.17)	(-0.97)	(1.40)	(1.79)	(17.59)	89.34%
2008	0.077	0.004	-0.002	0.003	0.006	0.007	(22.41)	(1.10)	(-0.55)	(1.07)	(2.00)	(1.54)	88.02%
2009	0.003	0.038	-0.002	0.004	0.006	0.053	(1.55)	(15.99)	(-0.87)	(1.85)	(2.42)	(22.33)	89.34%
2010	0.002	0.035	-0.002	0.002	0.003	0.038	(1.39)	(16.80)	(-0.69)	(1.02)	(1.53)	(15.83)	89.54%
2011	0.002	0.043	-0.002	0.002	0.003	0.050	(1.27)	(17.71)	(-0.84)	(1.04)	(1.50)	(18.56)	89.84%
2012	0.002	0.045	-0.003	0.002	0.003	0.039	(1.14)	(20.34)	(-1.05)	(1.20)	(1.54)	(17.30)	90.29%

Figure A1: **DY Tree.** For a given trading day, private information arrives with probability α. When there is no private information, buys and sells are Poisson with intensity ϵ_B and ϵ_S. Private information is good news with probability δ. The expected number of buys (sells) increases by μ in case of good (bad) news. Non-information related order flow shocks arrive with probability θ. In the event of an order flow shock, buys and sells increase by δ_b and δ_s respectively.

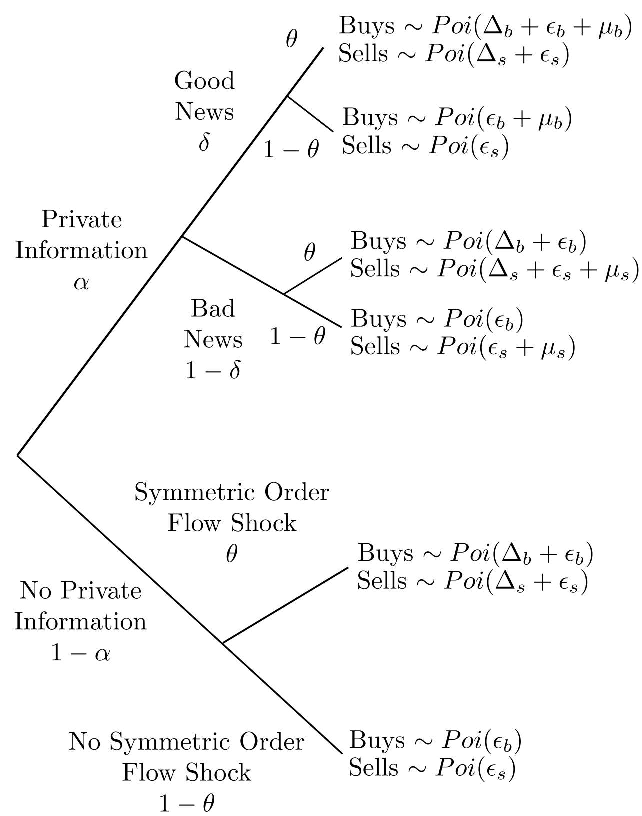


Figure A2: **XOM DY.** This figure compares the real and simulated data for XOM in 1993 and in 2012 using the DY model. In Panels A and B, the real data are marked as +. The real data are shaded according to the $CPIE_{DY}$, with darker markers (+ magenta) representing high and lighter markers (+ cyan) low $CPIE$s. The simulated data points are represented by transparent dots, such that high probability states appear as a dense, dark "cloud" of points, and low probability states appear as a light "cloud" of points. The DY model extends the three states of the PIN model corresponding to no news, good news, and bad news with three additional states with higher order flows due to non-information symmetric order flow shocks.

(a) XOM 1993 (b) XOM 2012

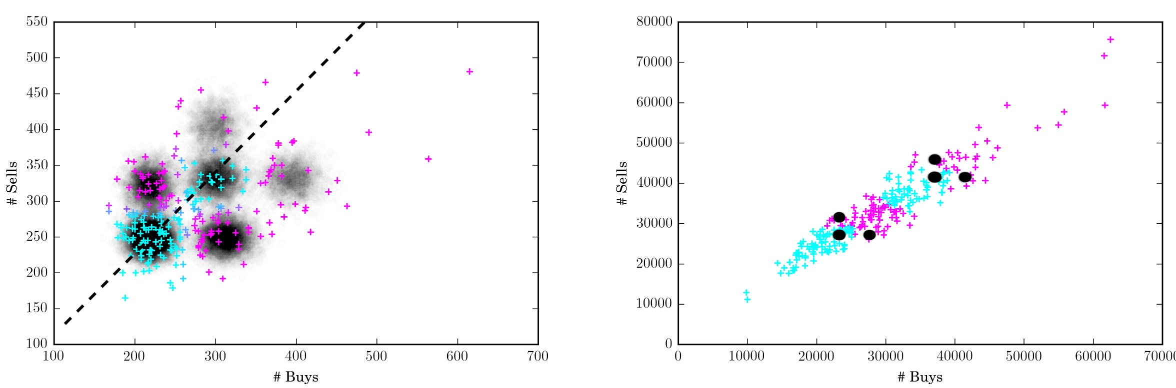


Figure A3: **Breakdown of the DY Model.** This figure shows the distribution of the percent of days where the total likelihood, given the model parameters and observed order flow data is less than 10^{-10}—days, according to the model, with near-zero probability of occurring. The solid black line represents the median stock, and the dotted lines represent the 5, 25, 75, and 95 percentiles.

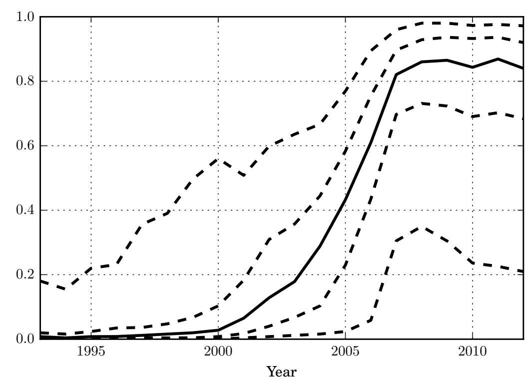
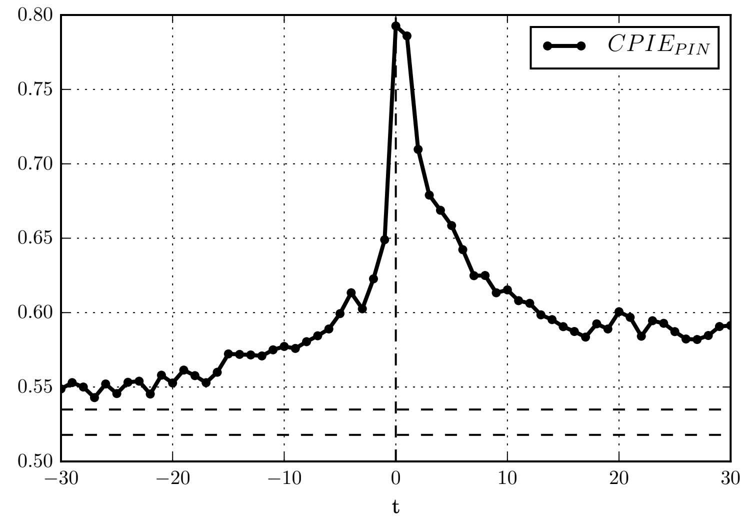
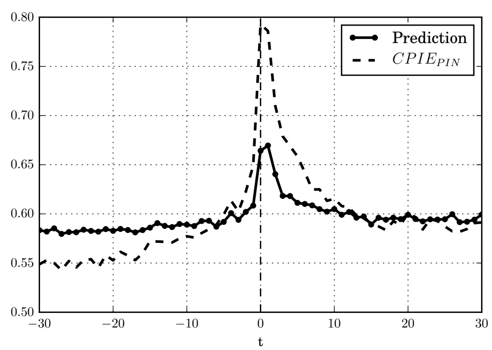


Figure A4: **M&A Targets - PIN.** Panel A shows the average $CPIE_{PIN}$ for the PIN model in event time surrounding M&A announcements in target stocks. Panels B and C compare the average $CPIE_{PIN}$ with the $CPIE_{PIN}$ predicted with either the absolute order imbalance or turnover, respectively. To obtain the predictions, we run regressions of daily $CPIE_{PIN}$ on $|B-S|$ or $turn$, and their respective squared terms.

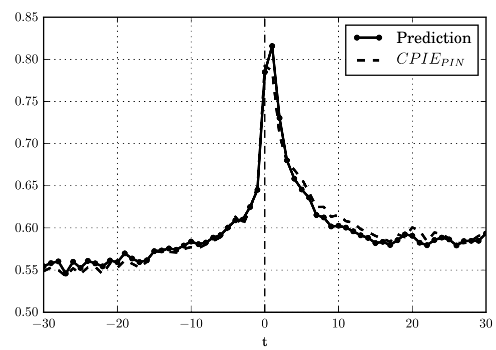
(a) $CPIE_{PIN}$



(b) Prediction using $|B-S|$ and $|B-S|^2$



(c) Prediction using $turn$ and $turn^2$



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